Do Mutual Funds Outperform During Recessions? International (Counter-) Evidence

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Abstract

Recent academic research documents that (i) U.S. equity mutual funds have a systematically better performance during periods of economic downturn and (ii) investors are willing to pay high fees for funds that provide recession insurance. In this paper, we test these hypotheses out-of-sample using international mutual fund data from 16 different countries. Surprisingly, we obtain contrary results: Based on our worldwide sample mutual funds underperform by a statistically significant -0.4% during months of economic downturn and funds with high recession alphas charge low fees to investors. We provide evidence that recession underperformance can be explained by fund managers' forced excessive trading.

JEL Classification: F30; G01; G11; G15; G23

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1 Introduction

Over the past few decades, the mutual fund industry worldwide has flourished as an investment vehicle for both retail and institutional investors. The number of mutual funds worldwide increased from about 50,200 in 1998 to about 73,000 in 2012. The assets managed grew from 9.6 trillion USD at the end of 1998 to 26.84 trillion USD at the end of 2012, with about 40% of the assets invested into equity mutual funds (see Investment Company Factobook (2013)).

While the global fund industry has gained importance as a whole, academic studies on the performance of mutual funds have mainly focused on the U.S. market. Among others, Malkiel (1995), Jensen (1968), and Fama and French (2010) show that actively managed U.S. equity mutual funds in general underperform the market, net of fees. However, unconditional performance measures may understate the value added by active mutual fund managers: Moskowitz (2000), Staal (2006) and Kosowski (2011) document that U.S. equity mutual fund managers perform significantly better during economic downturns than during economic upturns. So far, this empirical finding is not well understood: Is there a systematic outperformance of mutual funds during recessions? If so, do mutual fund managers outperform in recessions because they want to outperform or is it simply easier for them to outperform?

A first attempt to rationalize the answers to this puzzle is proposed by Glode (2011). He develops a model in which a fund manager can generate state-specific active fund returns. These conditional returns come at a disutility to the manager, as they require an effort to generate and their pay-off will be highest in states in which investors are willing to pay more for these returns. Hence, a fund manager exercises more effort in generating abnormal returns in times when the economy is performing badly because then the investors' marginal utility of consumption is large and investors are willing to pay high fund fees for this insurance. As a result, the theoretical model predicts that (i) on aggregate mutual funds outperform during economic downturns and (ii) cross-sectionally that – relative to other funds – funds with poor unconditional performance can charge high fees to investors because they earn abnormal state-specific returns during recessions.¹

This paper empirically tests hypotheses (i) and (ii) using a worldwide dataset of equity mutual funds in 16 different countries for the sample period from 1980 to 2010.² To determine whether a country is in a bad economic condition, we use the recession indicators from the National Bureau of Economic Research (NBER) for the USA and recession indicators from the Economic Cycle Research Institute (ECRI) for the 15 remaining countries.

Our analysis reveals the following surprising results. First, we do not find evidence that mutual funds outperform during recessions. To the contrary, based on results of our pooled worldwide sample, mutual funds *underperform* by a statistically significant -0.4% in the months of economic downturns.³ On a country level, we find that (depending on the respective regression model) mutual funds underperform during recessions in 12 (15) of the 16 countries with the underperformance being statistically significant at the 1% level. During our sample period from 1980 to 2010, we also find an underperformance of mutual funds during recessions within the U.S. This result is in contrast to earlier empirical findings of the literature. We reveal that the reason for the differences is the longer sample period applied in our study; we obtain similar results as, e.g., in Glode (2008) when we restrict the sample to the period from 1980 to 2005.

Second, we do not find that funds with poor unconditional performance can charge high fees to investors because they offer recession insurance in the form of high statespecific returns. Based on results of the pooled worldwide sample, we show that mutual funds in the quintile with the lowest unconditional performance charge the highest fees but

¹Glode (2011) and Glode (2008) - in an earlier working paper version - give empirical support that U.S. mutual funds indeed have a systematically better performance when the economy is in a bad state and that - relative to other funds - funds with poor unconditional performance tend to charge high fees and generate highly countercyclical returns.

²The sample period for each country begins with the availability of country-specific mutual fund data in the Morningstar database.

³Recession performance of funds is based on the Carhart (1997) four-factor model based on countryspecific market and accounting information. Our results remain stable if we compute alphas using alternative factor models or use regional market and accounting information (see Table 6).

also display the *lowest* recession performance. The difference in recession alphas between funds with the best unconditional performance and the worst unconditional performance is 1.07% per month and statistically significant at the 1% level.

We conduct robustness checks and additional analyses to shed light on the empirically documented underperformance of mutual funds during recessions. Our results remain stable when we compute fund alphas based on alternative asset pricing risk factors and use recession indicators for each country obtained from the OECD (instead of NBER and ECRI). We also find that recession underperformance is not specific to the fund's investment style. In addition, we investigate whether investors react with high future inflows into funds with high recession performance. We do not find evidence that funds who show superior state-specific performance receive higher future inflows when controlling for unconditional performance.

How can one explain the *negative* performance of funds during recessions? A potential explanation is that mutual fund managers are more active during recessions - however, this does not lead to state-specific outperformance, but in turn results in higher trading costs and finally worsens recession performance. Increased active management of mutual funds during recessions can occur as a result of two channels: First, fund managers may voluntarily deviate more strongly from the benchmark in order to outperform competitors based on compensation incentives (see, e.g., Massa and Patgiri (2009) and Huang, Sialm, and Zhang (2011)) and employment risk (Kempf, Ruenzi, and Thiele (2009)). Second, fund managers may be *forced* to be more active because investors tend to redeem their money from equity markets during periods of economic downturns (see, e.g., Longstaff (2004)). Hence, mutual funds are more likely to be engaged in asset fire sales in recessions than in non-recessions and have to liquidate their assets in a bad market environment characterized by high illiquidity and high average trading costs (Coval and Stafford (2007)). Consistent with these ideas our empirical results indicate that aggregate flows into mutual funds are negative (positive) during recessions (non-recessions) and that the average tracking error (i.e., a fund's deviation from its benchmark) of mutual funds

is higher in recessions than in non-recessions. Turning to performance implications, we find that during times of economic downturn, high tracking error funds underperform low tracking error funds by statistically significant -0.73% per month. This spread is magnified to -1.34% (-1.57%, -1.95%) when we restrict our analysis to months with aggregate fund flows below zero (below -0.01, below -0.02). When performing a cross-country regression of a country's average mutual fund recession performance on the average mutual fund recession tracking error (TE) as well as average aggregate recession fund flows (FF), we find that both TE and FF have positive (negative) explanatory power at the 10% significance level.

Finally, as an additional out-of-sample check, we investigate the recession performance of hedge funds using data from the TASS database in the period from 1994 to 2012. As in the case of mutual funds, we find compelling empirical evidence that hedge funds, on average, underperform during months of economic downturn. This state-specific underperformance is also observed across the majority of different hedge fund investment styles.

The rest of the paper is organized as follows. Section 2 gives an overview of the related literature. In Section 3 we describe our dataset and explain the methodology of our analysis. Section 4 provides the main empirical results of our study. Finally, in Section 5, we conclude.

2 Literature Review

Our study is related to two strands of literature. First, we contribute to the literature investigating performance measurement of mutual funds in an international context. Ferreira, Keswani, Miguel, and Ramos (2013) investigate the determinants of the performance of equity mutual funds in 27 countries. They document that in most of the countries actively managed funds underperform passive investment strategies. Keswani, Ferreira, Miguel, and Ramos (2014) find significant performance persistence of equity mutual funds around the world and show that performance persistence is related to differences in mutual fund industry development across countries. Cremers, Ferreira, Matos, and Starks (2015) show that actively managed funds in many countries choose portfolios that track their stated benchmark index closely. This degree of 'closet indexing' of active funds is positively associated with fees as well as negatively related to performance and exists less in countries with a higher market share of passive index funds. Finally, Breloer, Scholz, and Wilkens (2014) find that a majority of international equity mutual funds exhibit significant exposure to country/sector momentum indicating that these factors matter for risk-adjusted fund performance evaluation.

Second, our paper contributes to the literature of time-varying performance measurement of mutual funds. Moskowitz (2000), Staal (2006), and Kosowski (2011) document that risk-adjusted performance of U.S. mutual funds is negatively correlated with the business cycle and that mutual fund alphas are 1-3.5% p.a. higher in recessions than in expansions. Lynch and Souza (2012) and Badrinath and Gubellini (2012) document that this counter-cyclical outperformance depends on the fund's specific investment style and that for many fund styles, conditional outperformance switches from counter-cyclical to pro-cyclical over time.

Glode (2011) is the first to rationalize the empirical finding of counter-cyclical outperformance of mutual funds in a theoretical framework and shows that previous unconditional performance measures seem to be misspecified. In his model, a skilled, active fund manager is able to generate returns that depend on the state of the economy. Assuming rational investors, the fund manager will generate outperformance during economic downturns as the pay-off will be highest in states in which investors are willing to pay more for these returns. Moreover, mutual funds that perform well during recessions can charge higher fees as they provide an insurance for investors when the economy is in a bad state. Hence, in this setup, mutual fund investing and negative unconditional expected fund returns can simultaneously arise.⁴ Finally, in a related paper, Kacperczyk,

⁴Besides the consumption-based argument put forward by Glode (2011), there exist different alternative explanations to rationalize recession outperformance. Kothari, Shu, and Wysocki (2009) argue that

van Nieuwerburg, and Veldkamp (2013) provide evidence that fund manager abilities are time-varying and change with the business cycle. In particular, skilled managers successfully that pick stocks well in expansions also time the market well in recessions. Fund managers who exhibit this time-varying skill outperform the market by 50-90 basis points per year.

In this paper, we extend the literature by investigating the time-varying performance of domestic equity mutual funds using a worldwide sample of 16 different countries. In particular, this paper is the first to empirically test on a global scale whether (i) mutual funds outperform during recessions and (ii) whether funds with poor unconditional performance can charge high fees to investors because they earn state-specific abnormal returns during recessions.

3 Data and Methodology

From the Morningstar mutual fund database, we retrieve data on all actively-managed open-end equity mutual funds domiciled in Australia, Austria, Canada, China, Denmark, France, Germany, India, Italy, Japan, Mexico, Norway, South Africa, South Korea, Spain, Sweden, Switzerland, Taiwan, United Kingdom, and the USA.⁵ The sample time period ranges from 1980 until 2010 with some countries having shorter time periods due to data availability. We narrow the data sample to only domestic equity mutual funds.⁶

As the mutual fund data is reported on a share class level, the multiple share classes are aggregated based on the fund's total net assets (TNA).⁷ To have a sufficient number

corporate managers delay disclosure of bad news relative to good news. Hence, during recessions, experienced investors, such as fund managers, could have an informational advantage over unexperienced retail investors and therefore earn higher returns. Other potential explanations include time-varying trading and liquidity costs during recessions and expansions (see e.g., Kosowski (2011)). Finally, time variation in risk exposures (e.g., time-varying market betas of mutual funds in recessions and expansions) without any active portfolio rebalancing can potentially explain outperformance during economic downturns.

⁵The selection of countries includes all continents and uses the United Nations Development Index (HDI) to select countries with higher financial education.

⁶Funds are identified by their classifications 'Global Category', 'Morningstar Category' and 'Investment Area'.

⁷As a robustness check, we perform our empirical analysis just based on the share class with the maximum TNA. Our main results remain unchanged.

of observations in our regression analysis in Section 4, we delete countries that have less than 30 funds, which removes Austria from our sample. Furthermore, we drop all funds from the sample that do not go through at least one recession and have less than 12 months of observations.

To determine whether a country is in a poor or good economic condition, we use the recession indicators from the National Bureau of Economic Research (NBER) for the USA and the recession indicators from the Economic Cycle Research Institute (ECRI) for the remaining countries.⁸ We delete all countries that do not go through at least one recession or spent less than 5% of their sample time in a recession. This restriction removes Australia, China and India from our sample. All the other countries spend about 10% to 40% of their sample time in recessions.

We display summary statistics and data availability for all countries in Table 1. The average time a country spends in a recession is displayed in Panel A of Figure 1.

[Insert Table 1 about here]

[Insert Figure 1 about here]

Table 1 shows that the highest number of domestic equity funds are located in the USA (3,692), followed by Japan (811), and South Korea (524). U.S. equity funds also have the highest average fund TNA with a mean value of 553m USD.

From Panel A of Figure 1, we observe that Italy (41.9% of the time) and UK (41.6% of the time) spend most of their sample time in a recession. The least time in recessions is spent by South Korea (11.4%) and the USA (15.8%).

In Panel B of Figure 1, we take a look at the percentage of countries that are in a recession at the same point in time. The first subplot shows the development of recession clustering in all countries over time. We find that the number of worldwide recessions is

⁸ECRI does not provide recession indicators for Norway and Denmark. Instead we retrieve recession indicators for those countries from the OECD business cycle measure as in Christoffersen (2000) and Steigum (2004). We also use the OECD business cycle indicators for all countries in our sample as a robustness test in Section 4.3.1.

high during the beginning of the 1980s, 1990s, 2000s, and in 2009, which can be related to worldwide economic downturns (oil price shock and restrictive monetary policy of the FED in the 1980s, banking crises at the beginning of the 1990s, the burst dot-com bubble in the beginning of the 2000s, and the financial crisis in 2009 following the collapse of the U.S. mortgage market). The subplots North America, Europe, Asia and rest of the world show the percentage of countries that are in a recession at the same point in time in a particular geographical region. Hence, one can infer that there are time periods in which recession periods in the different geographical regions do not overlap.

To get an impression of differences in unconditional mutual fund performance across countries, we provide average monthly fund returns for all countries in Figure 2.

[Insert Figure 2 about here]

It shows that the highest average returns in percent per month are found in South Africa (1.18%) and Sweden (0.97%), whereas the lowest returns are found in Italy (-0.37\%) and Japan (-0.21\%). The average monthly return over all funds and all countries is 0.48%.

Performance Evaluation. In our empirical analysis in Section 4, we evaluate conditional mutual fund performance (in recessions) using the Carhart (1997) four factor model. We differentiate between two specifications.

Specification (1) estimates the Carhart (1997) model including a business cycle dummy (BC) variable to account for recession performance

$$r_{it} = \alpha_{it} + BC_t + \beta_1 RMRF_t + \beta_2 SMB_t + \beta_3 HML_t + \beta_4 MOM_t.$$
(1)

In specification (2), we additionally interact the asset pricing risk factors with the business cycle variable

$$r_{it} = \alpha_{it} + BC_t + \beta_1 RMRF_t + \beta_2 RMRF_t \times BC_t + \beta_3 SMB_t + \beta_4 SMB_t \times BC_t + \beta_5 HML_t + \beta_6 HML_t \times BC_t + \beta_7 MOM_t + \beta_8 MOM_t \times BC_t,$$
(2)

where r_{it} is the monthly fund return in excess of the risk-free rate, RMRF is the market factor, SMB is the size factor, HML is the value factor, and MOM is the momentum factor.

We use monthly factor returns and approximations for the risk-free rate from two different sources. Data for the USA is taken from Kenneth French's webpage; for the remaining countries we obtain data from the webpage of Sandy Lai.⁹ As a robustness check in Section 4.3.1, we also verify our results using the individual factor returns obtained from Andrea Frazzini's webpage¹⁰ and Stefano Marmi's webpage¹¹, as well as the international regional factor returns from the webpage of Kenneth French.¹² Kenneth French provides regional risk-free interest rates, market, size, and book-to-market factors for Europe, North America, Japan and Asia. We assign our sample countries to the different factors based on their geographical location. We also test the stability of our results by calculating own individual factor returns on the basis of domestic total return indices retrieved from Datastream.¹³ The start dates for mutual fund data in Morningstar and the different factor returns for all individual countries in our sample are displayed in Table 1.

⁹The dataset contains the four Carhart (1997) factors for all countries in our sample usually beginning in the 1980s. Data can be obtained from http://www.sandylai-research.com/html/research___data.html and is also described in Eun, Lai, de Roon, Zhang (2010) and Hau and Lai (2013).

¹⁰Data Library: http://www.econ.yale.edu/~af227/data_library.htm

¹¹Data Library: http://homepage.sns.it/marmi/Data_Library.html

¹²Data Library: http://mba.tuck.dartmouth.edu/pages/faculty/ken.french/data_library.html

¹³We approximate the market portfolio using the broadest equity market index of the country and the risk-free rate by a domestic deposit rate. The SMB factor is calculated as the difference between the monthly returns of the small cap and large cap equity indices, and the HML factor is computed as the difference between the domestic value and growth indices. A disadvantage of using the domestic indices from Datastream is their short availability and that these indices are not appropriate for constructing the momentum factor.

4 Empirical Results

4.1 Aggregate Mutual Fund Performance During Recessions

4.1.1 Pooled Sample

We investigate the conditional performance of equity mutual funds during recessions in a worldwide sample of 16 different countries. As described above, we apply the Carhart (1997) four factor model to investigate risk-adjusted performance. If not otherwise specified, the return factors for the USA are obtained from Kenneth French's webpage and for the remaining countries from Sandy Lai's webpage. To identify recession periods in the respective country, we use the NBER indicators for the USA and the ECRI indicators for the remaining countries. Recession indicators are indicated by the *BC* variable, a dummy variable that takes on the value one if the country is in a recession and zero if not. Table 2 provides results of different panel regressions with fixed effects on the fund level for the worldwide sample of all funds in 16 countries.

[Insert Table 2 about here]

Regression (1) of Table 2 documents the results of a regression of monthly excess returns on the four Carhart (1997) factors. In line with previous research (see, e.g., Ferreira, Keswani, Miguel, and Ramos (2012), Ferreira, Keswani, Miguel, and Ramos (2013), and Cremers, Ferreira, Matos, and Starks (2015)), we document negative unconditional performance of mutual funds with an alpha of -0.0229% per month, which is statistically significantly different from zero at the one percent level. In addition, our results indicate that mutual funds display positive factor loadings on the market, SMB, and HML as well as a negative loading on momentum (MOM).

Regression (2) shows the results when we extend our model with the BC recession dummy. BC has a coefficient estimate of -0.402 and is statistically significant at the one percent level. Hence, on our worldwide sample, mutual funds *underperform* during times of recessions by -0.402% per month based on the Carhart (1997) four factor model. Regressions (3) and (4) re-estimate specification (1) in a subsample of recession and nonrecession months, respectively. In line with the results of regression (2), we find that the alpha is negative during recession months and slightly positive in non-recession months.

Finally, in regression (5), we account for time-varying factor sensitivities by additionally including interaction terms of the business cycle dummy with the Carhart (1997) factors. The result of negative fund performance during recessions remains unchanged: mutual funds underperform during recession months by statistically significant -0.480%. In unreported tests we document that our results also remain unchanged if we evaluate mutual fund performance using the Fama and French (1993) three-factor model. In each case, we obtain a significantly negative impact of the recession dummy on aggregate mutual fund performance.

4.1.2 Country-Specific Analysis

We proceed to analyze country-specific risk-adjusted performance of mutual funds in recessions. Table 3 repeats regressions (2) and (5) of Table 2 separately for the 16 different countries of our sample.

[Insert Table 3 about here]

Panel A of Table 3 shows the results of country-specific regressions of excess returns on the Carhart (1997) factors and the BC recession dummy variable. Strikingly, we find that in 15 of the 16 countries the BC dummy has a negative impact on the performance of mutual funds and is statistically significantly different from zero at the one percent level. The most negative impact of recessions on the performance is found in South Africa (-2.181% per month), Switzerland (-1.391% per month), and Sweden (-1.209% per month). The only positive (but statistically insignificant) impact of BC is found in Germany with a tiny outperformance of 0.035% per month.

In Panel B, we redo the analysis of Panel A but also include the interaction terms of all Carhart (1997) factors with the *BC* variable. Our results remain qualitatively unchanged;

we find significant outperformance of mutual funds during recessions only in Germany (0.189% per month) and Spain (0.344% per month), while in 11 of the 16 countries *BC* is negative at the one percent significance level. The most negative impact of recessions on the performance of mutual funds is found in South Africa (-2.274\% per month), Sweden (-1.319\% per month) and Norway (-0.988\% per month). In untabulated results we also run regression specifications (3) and (4) of Table 2 for all individual countries in our sample. In line with our previous results, the alpha during recession periods is worse than the alpha in non-recessions in 12 of the 16 countries.

To summarize, we do not find evidence that mutual funds outperform during periods of economic downturn; instead we find strong evidence that mutual funds *underperform* during recessions. This result not only holds on our pooled worldwide sample (as shown in Section 4.1) but is also valid for the majority of individual countries.

Revisiting Fund Performance During Recessions in the USA. Panel A of Table 3 indicates that for the USA, BC is significantly negative at the one percent significance level. When interacting all Carhart (1997) factors with the BC variable in Panel B, we obtain a slightly positive impact of the recession dummy; however, the effect is not significantly different from zero. Both findings are not in line with the theoretical model of Glode (2011) and the empirical results displayed in Glode (2008).

To evaluate these differences in results, we replicate the sample of Glode (2008) using CRSP mutual fund data and identical data cleaning procedures for the USA (see Glode (2008), Table 1) in the time period from 1980 - 2005 (Glode (2008)'s sample period) and from 1980 - 2010 (our sample period). Results are displayed in Table 4.

[Insert Table 4 about here]

Regression (1) shows regression (3) of Table 3 in Glode (2008) with CRSP mutual fund data from 1980 - 2005. The impact of the recession dummy is positive and statistically significant at the one percent significance level indicating an average mutual funds' outperformance of 0.414% per month. In regressions (2) and (3), we replicate the empirical results of Glode (2008) for the identical time period (1980 - 2005) using data from CRSP and Morningstar, respectively. In both cases, we obtain similar results: The coefficient estimate for BC is positive and statistically significant at the one percent significance level. Then, in regressions (4) and (5), we expand the sample period until the year 2010, again using data from CRSP and Morningstar, respectively. Surprisingly, the extended sample period now delivers different results: For both datasets, we do not obtain a significant positive correlation between recessions and average mutual fund performance – instead, the relationship is insignificant for both datasets.

Our results show that differences in results for the USA between our study and Glode (2008) can be attributed to the extended sample period from 1980 - 2010. Instead of finding support in favour of a positive relationship, we fail to do so and document no significant relationship between recessions and average mutual fund performance in the USA.¹⁴

4.2 Cross-Sectional Implications: Unconditional Performance, Fund Fees, and State Dependence

Glode (2011)'s theoretical model predicts that – relative to other funds – funds with poor unconditional performance can charge high fees because they exhibit more state dependence in realized performance, i.e., earn abnormal state-specific returns during recessions. Hence, we investigate whether state dependence in aggregate fund performance is mostly driven by funds with poor unconditional performance and whether insurance against bad states of the economy might partially explain the survival of some of the most poorly performing funds.¹⁵

Since Morningstar does not contain historical fund fee data, we use the last fund fee

 $^{^{14}}$ When we do not include interaction terms of the *BC* dummy with the Carhart (1997) factors, we find a negative relationship between recessions and mutual fund performance for the USA, see Panel A of Table 3.

¹⁵Glode (2011) provides empirical evidence for this relationship in the USA for the time period from 1980 to 2005. He finds that the (unconditionally) worst performing funds charge the highest fees and display the best recession performance.

observation from the Morningstar Mutual Fund Database as of September, 2012 in our empirical analysis. It is important to note that this does not create a survivorship bias as the last fund fee observation is also stored for dead funds.

As in Sirri and Tufano (1998), Barber, Odean, and Zheng (2005) and Glode (2011), we compute the total fund fees as "the expense ratio plus the up-front load amortized over a seven-year holding period (which is the average holding period for equity mutual funds)" (Sirri and Tufano (1998)). If there is a fee schedule for the up-front load, we use the starting fee for the lowest investment amount. We set negative expense ratios to zero.¹⁶ Figure 3 plots the average total fund fees (in % of a fund's TNA) per country.

[Insert Figure 3 about here]

The highest average total fees in percent are found in Italy (2.65%) and France (2.58%) whereas the lowest average fees are found in Sweden (1.27%) and South Korea (1.38%).

In order to test the cross-sectional prediction on our worldwide dataset, we sort funds into quintile portfolios based on their average monthly unconditional performance according to the four-factor Carhart (1997) model. Then, we compute the average annual expense ratio, the average total annual fee and average monthly recession performance of the respective quintiles. To measure fund-specific recession performance, we use the BCcoefficient in specification (1) estimated over the whole return series of each individual fund. Our results remain qualitatively the same if we use specification (2) instead. Panel A of Table 5 shows the results for the pooled worldwide sample.

[Insert Table 5 about here]

Consistent with previous results in the literature (see, e.g., Malkiel (1995) and Carhart (1997)), we find that there is a negative relationship between a fund's unconditional performance and a fund's fee. The quintile with the best (worst) unconditional performance charges average expense ratios of 1.47% p.a. (1.80% p.a.) as well as average total expenses of 2.39% p.a. (1.89% p.a.). Hence, the quintile with the lowest unconditional

 $^{^{16}\}mathrm{Fund}$ fees are available for all countries in our sample except for Canada.

performance charges expense ratios (total fund fees) of 0.33% p.a. (0.50% p.a.) higher than the quintile of funds with the highest unconditional performance. The differences are statistical significant at the one percent level, respectively. More importantly for our research question, we find that funds in the quintile with the *lowest* unconditional performance also display the *lowest* recession performance. The difference in recession alphas between funds with the best unconditional performance and the worst unconditional performance is 1.07% per month and is statistically significant at the 1% level. Hence, we do not find evidence that funds with poor unconditional performance can charge high fees because they earn superior state-specific returns during recessions.

Panel B of Table 5 repeats the same analysis for the individual countries in our sample. We only report differences in the average annual expense ratio, the average total annual fee and average monthly recession performance between quintile portfolio 5 (funds with the best unconditional performance) and quintile portfolio 1 (funds with the worst unconditional performance). In all countries we find that funds with low unconditional performance also display *low* recession performance. The differences in recession alphas between funds with the best unconditional performance and the worst unconditional performance are positive and statistically significant at the 1% in all countries. They vary between 0.46% per month in Italy and 1.56% per month in the UK.

In summary, our results indicate that there is a negative relationship between a mutual fund's unconditional performance and its expense ratio (total fee). However, we do not find evidence that funds with poor unconditional performance can charge high fees because they earn abnormal state-specific returns. To the contrary, funds with poor unconditional performance also exhibit poor recession performance.¹⁷

¹⁷In Table A.1 of the Appendix, we investigate the relationship between a fund's recession performance and its fund fee. Consistent with the results of Table 5 we find that funds with superior recession performance tend to charge lower expense ratios and lower total fund fees.

4.3 Robustness and Additional Analyses

4.3.1 Risk Factors and Business Cycles

In this section we perform different robustness checks to confirm our main result of mutual funds' negative performance in recessions. In particular, we show that our results are stable if we use different asset pricing risk factors and alternative business cycle measures. Table 6 reports the results using the pooled worldwide sample.

[Insert Table 6 about here]

Risk Factors. Regression (1) repeats our baseline setup from Table 2 using the monthly factor returns from Sandy Lai. In regressions (2) and (3) we apply the asset pricing risk factors obtained from Andrea Frazzini and Stefano Marmi. In both cases, we obtain negative coefficient estimates for BC of -0.34 and -0.11, which are both statistically significant at the one percent level. As risk factors, regression (4) uses self-constructed individual factor returns which are based on domestic total return indices retrieved from Datastream. Our results are virtually unchanged. Finally, in regression (5), we verify our results using international regional factor returns from Kenneth French. Again, we obtain a negative coefficient estimate for BC of -0.33, which is statistically significant at the one percent level.

OECD business cycle measure. The business cycle measure is an important variable that determines the number of months a country spends in a recession. To gain robustness in our main findings, we repeat our investigation with an alternative business cycle measure obtained from the Organisation for Economic Co-operation and Development (OECD). Recession and expansion periods are signified by their deviation from a growth trend. The main reference for this trend is the industrial production of the respective countries.¹⁸ Regression (6) shows the results when we categorize recessions based on the

 $^{^{18}{\}rm The}$ OECD indicator can be retrieved from the OECD database: http://www.oecd.org/std/leading-indicators/.

OECD business cylcle measure: Mutual funds show a statistically significant state-specific underperformance of -0.28% per month.

To conclude, we find that our main result of mutual funds' negative recession performance is stable if we use alternative asset pricing risk factors and business cycle measures. Detailed results of the robustness checks for the individual countries in our study are shown in Table A.2.

4.3.2 Fund Style Analysis

Badrinath and Gubellini (2012) find for the USA that funds with different investment styles display different state-specific performance. In particular, they argue that managers of small-cap and mid-cap growth equity funds are able to deliver recession outperformance but managers of value funds are not. We also test whether alternative styles of mutual funds differ in their recession performance based on our pooled worldwide dataset.¹⁹ We run regression specification (2) for small/mid-cap/large as well as value/growth/income funds.

Table 7 shows that all fund styles display a negative recession performance, as indicated by a negative coefficient estimate for the BC variable. The most negative recession performance is shown by the income fund style (-0.524% per month) followed by the large (-0.513% per month) and mid-cap (-0.398% per month) fund style. Hence, we do not find that conditional performance during recessions can only be attributed to certain mutual fund styles. Instead, negative performance during economic downturns seems to be a general phenomenon.

4.3.3 Recession Performance and Fund Flows

Glode (2011)'s model assumes that investors are willing to pay for high returns when

¹⁹The identification of fund styles is based on the Morningstar fields 'Broad Category', 'Global Category', and 'Morningstar Category'.

the economy is in a bad state. If this is the case, one should see a positive impact of a fund's recession performance on future fund inflows controlling for a fund's unconditional performance.

Following Sirri and Tufano (1998), Guercio and Tkac (2002) and Ferreira, Keswani, Miguel, and Ramos (2013), we calculate fund i's flow in month t as

$$flow_{i,t} = \frac{TNA_{i,t} - TNA_{i,t-1} \cdot (1+r_{i,t})}{TNA_{i,t-1}},$$
(3)

where $TNA_{i,t}$ and $r_{i,t}$ denote the fund total net asset value and raw return of fund iin month t.²⁰

To test the idea that funds with a top performance during recessions receive additional inflows when controlling for unconditional performance, we perform portfolio doubls sorts. In a first step, in month t, we sort funds based on their unconditional Carhart (1997) four factor alphas over the past 60 months. Then, within each quintile, we sort funds based on their past recession alphas over the same time period. We analyze fund flows in month t + 1 for these 25 portfolios.²¹ Table 8 reports the results of the double sorts for the pooled worldwide sample.

[Insert Table 8 about here]

In line with the results of Sirri and Tufano (1998) and Ferreira, Keswani, Miguel, and Ramos (2013), we document a strong positive relationship between unconditional performance and future fund flows. Differences in monthly future fund flows between the unconditionally best and worst performing funds are positive in all (conditionallysorted) quintiles with spreads ranging from 0.46% to 3.70% with an average spread of 1.55%. In contrast to the idea that investors particularly value outperformance during recessions, we do not observe higher inflows to funds with high recession alphas. Our

 $^{^{20}}$ To ensure that extreme values do not drive our results, we truncate outliers above -100% and 500% and winsorize fund flows at the bottom and top one percent level of the distribution.

²¹The results remain qualitatively the same if we reverse the order of the double sorts or if we sort based on past raw returns.

results indicate that the differences in future fund flows between the conditionally best and worst performing funds are negative in four of the five (unconditionally-sorted) quintiles. The average spread between the conditionally best and worst performing funds amounts to -0.38%. Hence, our results based on portfolio double sorts do not support the idea that investors value a fund's recession performance in addition to a fund's unconditional performance when investing in mutual funds.

In Table A.3 of the Appendix, we also report results of univariate portfolio sorts based on unconditional performance and future fund flows as well as conditional performance and future fund flows. Again, we observe a positive, statistically significant relationship between unconditional Carhart (1997) four factor alphas and future fund flows. However, this does not prevail in the conditional performance sort: we do not find a statistically significant relationship between a fund's recession performance and future fund flows.

4.4 Explaining Mutual Fund Underperformance During Recessions

How can one explain the negative performance of mutual funds during recessions? In this section, we provide empirical evidence that is consistent with the notion that mutual fund managers are more active during recessions - however, more activeness does not lead to state-specific outperformance, but results in higher trading costs and underperformance during recessions.

To measure fund manager activity over the business cycle, we investigate mutual funds' *tracking errors* in recessions and non-recessions. We compute the tracking error of mutual fund i as the square root of the second moment of the difference between r_i and the main domestic stock market index return r_m in the respective country:

$$TE_i = \sqrt{E(r_i - r_m)^2}.$$
(4)

Panel A of Table 9 reports equal-weighted averages of mutual funds' tracking errors

in recessions and non-recessions on the pooled sample and individual countries.

[Insert Table 9 about here]

We empirically find that, overall, the average tracking error during recessions is 4.16, whereas it is 3.41 in non-recessions. The difference of 0.74 is statistically significant at the one percent level and also economically significant. In addition, we find that in 15 of the 16 countries, mutual funds show a tracking error that is higher during recessions than in non-recessions.

Why are mutual fund managers more active during recessions? We identify that increased active management of mutual fund managers during recessions can occur as a result of two channels. First, fund managers may voluntarily deviate more strongly from the benchmark. They may attempt to outperform competitors based on compensation incentives (see, e.g., Massa and Patgiri (2009) and Huang, Sialm, and Zhang (2011)) or employment risk (Kempf, Ruenzi, and Thiele (2009)). Second, fund managers may be forced to be more active because, during periods of economic downturns, investors tend to redeem their money from equity markets (see, e.g., Longstaff (2004)). Hence, mutual funds are more likely to be engaged in asset fire sales in recessions than in nonrecessions and must liquidate their assets in a bad market environment characterized by high illiquidity and high average trading costs (Coval and Stafford (2007)). We report the level of aggregate flows (computed as the mean of equal-weighted individual fund flows per month, averaged over the time series) into mutual funds for recessions and nonrecessions in Panel B of Table 9. We observe that aggregate flows into equity mutual funds are positive 1.61% in non-recessions, whereas they are negative -0.09% in recessions. The difference between aggregate flows in recessions and non-recessions based on the worldwide sample amounts to -1.70% and is statistically significant at the one percent level. In addition, we find that aggregate fund flows into mutual funds during recessions are negative in 10 of 16 countries – hence, it is likely that fund managers are forced to trade because investors redeem money during months of economic downturn.²²

²²Note that fund flows out of mutual funds frequently lead to an immediate trading pressure for fund

In order to investigate the relationship between fund manager activity and statespecific performance, we sort funds into quintiles based on their tracking error during recessions and investigate conditional alphas based on the Carhart (1997) four factor model. Table 10 reports the results.

[Insert Table 10 about here]

Our results indicate that mutual funds with a high tracking error perform worse than low tracking error mutual funds during recessions.²³ Conditional alphas are -0.73% lower for high tracking error funds than for low tracking error funds. Moreover, high tracking error funds underperform low tracking error funds in 10 of 16 countries during recessions.

To investigate whether recession underperformance of funds is magnified by the likelihood of forced trading, we restrict our analysis to months with aggregate fund flows below zero (below -1%, below -2%). Indeed, we observe that high tracking error funds underperform low tracking error funds by -1.34% (-1.57%, -1.95%) when we restrict our analysis to months with negative aggregate fund flows. Hence, our results are consistent with the notion that forced active trading of mutual funds plays an important role to explain mutual fund underperformance during recessions.

Finally, we seek to explain the determinants of mutual fund performance in a multivariate regression framework. Table 11 reports the results.

[Insert Table 11 about here]

In regressions (1) and (2) we regress country *i*'s average mutual fund recession performance in month t on country *i*'s average mutual fund recession tracking error (TE) as well as average aggregate recession fund flows (FF) in month t. We find that TE has a negative coefficient estimate of -0.1659 (significant at the 5% level) and FF has a positive coefficient estimate of 0.2623 (statistically significant at the 5% level). In regression (3)

managers, whereas inflows can be held as cash and invested within a longer-lasting time period.

 $^{^{23}}$ This result is in contrast to our results based on unconditional performance. Untabulated results show that the relationship between a fund's tracking error and unconditional performance is significantly positive (at the one percent level). This is in line with the empirical findings of Huij and Jeroen (2011).

we use both TE and FF as regressors – again, we find that both independent variables remain statistically significant at least at the 10% level. Finally, in regression (4), we add different country characteristics such as GDP per capita, the number of listed domestic stocks, mutual fund assets per GPD per capita, the number of mutual fund companies per country, and a country's home bias to our model.²⁴ Controlling for these additional variables, we still find that both TE and FF have a statistically significant impact on a country's average mutual fund recession performance.²⁵

To summarize, we provide empirical evidence that mutual fund underperformance during recessions is driven by excessive trading of mutual fund managers. During months of economic downturn funds with high tracking error underperform funds with low tracking error. The underperformance is magnified during months where fund managers are likely to be engaged in asset fire sales.

4.5 Hedge Fund Performance During Recessions

As an additional out-of-sample analysis, we investigate the recession performance of hedge funds. We obtain hedge fund returns and characteristics from the TASS database in the time period from 1994 until 2012. Summary statistics of the sample of hedge funds are shown in Table A.4 in the Appendix.

We perform regressions similar as in Table 2 with the pooled sample of hedge funds instead of mutual funds applying hedge fund specific risk factors. Panel A of Table 12 reports the results.²⁶ To control for specific hedge fund risk factors, we use the seven

²⁴We obtain data for GDP per capita, the number of listed domestic stocks, mutual fund assets per GPD per capita from Datastream and data for mutual fund assets per GPD per capita, the number of mutual fund companies per country from Morningstar. Data for a country's home bias is obtained from Chan, Covrig, and Ng (2005) which is based on all mutual fund holdings in a respective country in 1999/2000. It signifies how much the mutual fund holdings of country i in the domestic market deviate from its holdings in the world market portfolio. Higher relative domestic holdings indicate a more pronounced home bias.

²⁵We also find that mutual fund performance is lower in countries that show a stronger home bias and higher in countries with a larger number of mutual funds.

²⁶The hedge fund sample is divided into the following trading strategies: Fixed Income Arbitrage, Dedicated Short Bias, Emerging Markets, Equity Market Neutral, Event Driven, Fund of Funds, Global Macro, Long/Short Equity Hedge, Managed Futures, and Multi-Strategy.

factor model of Fung and Hsieh (2004). These seven factors comprise a market factor derived from the S&P 500 monthly total return as well as a size factor calculated from the spread between the monthly return of the Russell 2000 and the S&P 500 index. Furthermore, the monthly change in the 10-year treasury yield, and the monthly change in yields between Moody's Baa yield and the 10-year treasury constant maturity yield are contained as factors. In addition, three trend-following factors for foreign-exchange, commodities, and bonds are included.²⁷

[Insert Table 12 about here]

Regression (1) documents the results of a unconditional regression of monthly excess returns on the seven hedge fund risk factors. In line with the results of Fung and Hsieh (2004), hedge funds show a significantly positive unconditional alpha. Regression (2) shows the results when we append our model with the BC recession dummy. BC has a coefficient estimate of -0.00571 and is statistically significant at the one percent level. Regressions (3) and (4) re-estimate the seven factor model of Fung and Hsieh (2004) in a subsample of recession and non-recession months, respectively. In line with the results of regression (2), we find that the alpha is negative during recession months and positive in non-recession months. Finally, in regression (5), we account for time-varying factor sensitivities by additionally including interaction terms of the business cycle dummy with the hedge fund risk factors. Again, the BC recession dummy remains negative and statistically significant at the one percent level.

Regressions (1) to (10) in Panel B of Table 12 report the recession performance for the subsamples of different hedge fund strategies. We find that seven out of ten hedge fund trading strategies display a negative BC recession dummy – significant at the one percent significance level. Hence, our results also confirm recession underperformance for a wide range of different hedge fund strategies.

 $^{27}The trend-following factors are obtained from David Hsieh's webpage: https://faculty.fuqua.duke.edu/~dah7/HFRFData.htm.$

5 Conclusion

Recent academic research (e.g., Moskowitz (2000), Kosowski (2011), and Glode (2011)) shows that U.S. mutual funds outperform during times of economic downturn and that funds with high performance during recessions can charge high fees to investors. Our paper is the first to test these hypotheses empirically applying a worldwide sample of domestic equity mutual funds.

Using mutual fund data from 16 different countries in the sample period from 1980 to 2010 and applying recession indicators from the NBER, ECRI and OECD as measures of economic downturn as well as factor data from six different factor data sets, our analysis reveals the following results: First, we are not able to find evidence that mutual funds outperform during recessions; in contrast, our results indicate that mutual funds underperform by statistically significant -0.4% in months of economic downturn worldwide. Second, we show that mutual funds in the quintile with lowest unconditional performance charge the highest fees but also display the lowest recession performance. The difference in recession alphas between funds with the best unconditional performance and the worst unconditional performance is 1.07% per month and statistically significant at the 1% level.

We provide empirical evidence that increased activity of mutual fund managers can explain the negative performance of funds during recessions. For this purpose, we document that the average tracking error of mutual funds is higher in recessions than in nonrecessions and that, during times of economic downturn, high tracking error funds underperform low tracking error funds by statistically significant -0.73% per month. Moreover, we find that this underperformance is magnified for funds that are likely engaged in asset fire sales and have to liquidate their assets in a bad market environment characterized by high illiquidity and high average trading costs.

Appendix

Table A.1: Recession Performance and Fund Fees

This table shows the total fund fees (Panel A) and fund expense ratios (Panel B) sorted according to their recession fund performance in 14 countries and for the pooled sample. The total fund fees are calculated as described in Section 3 and presented in %. Canada and Japan are missing from the list of countries, as there not enough observations of the fund fees. Statistical significance at the ten, five and one-percent level is indicated by *,**, and ***, respectively. T-statistics are displayed in parentheses.

Panel A: Total Fund Fees sorted on Recession Performance

I and A. I	lotari	unu		oricu	011 10		entormance
	$\mathbf{Q1}$	$\mathbf{Q2}$	$\mathbf{Q3}$	$\mathbf{Q4}$	$\mathbf{Q5}$	Q5-Q1	Q5-Q1 t-value
ALL	2.05	2.06	2.14	2.05	1.97	-0.08	-1.14
Denmark	1.59	1.83	1.72	1.13	1.32	-0.27	-0.78
France	2.62	2.91	2.56	3.37	2.46	-0.15	-0.59
Germany	2.63	2.35	2.24	2.36	2.90	0.27	0.93
Italy	2.56	3.32	2.67	2.81	2.92	0.36	1.07
Mexiko	2.38	2.30	2.07	2.12	2.48	0.09	0.13
Norway	1.58	1.66	1.61	1.47	1.38	-0.21	-0.74
South Africa	1.90	1.84	1.81	2.03	1.92	0.01	0.10
South Korea	1.53	1.26	1.54	1.39	1.34	-0.19*	-1.93
Spain	2.19	2.16	2.12	1.82	1.78	-0.41*	-2.50
Sweden	1.99	0.88	1.02	1.31	1.06	-0.92*	-2.17
Switzerland	1.78	1.94	1.44	1.50	2.04	0.26	1.09
Taiwan	2.03	1.99	1.87	1.95	2.00	-0.03	-0.53
UK	2.08	2.03	1.65	1.77	1.84	-0.24**	-2.20
\mathbf{USA}	2.28	2.04	2.09	2.04	2.15	-0.13	-0.88

Panel B: Expense Ratio sorted on Recession Performance

	01	00	01	0.4	05	05 01	
	$\mathbf{Q1}$	$\mathbf{Q2}$	$\mathbf{Q3}$	$\mathbf{Q4}$	$\mathbf{Q5}$	Q5-Q1	Q5-Q1 t-value
ALL	1.65	1.29	1.32	1.31	1.44	-0.21***	-3.97
Denmark	1.35	1.59	1.47	0.95	1.12	-0.23	-0.67
France	2.16	2.54	2.16	2.98	2.06	-0.10	-0.41
Germany	2.02	1.65	1.57	1.80	2.11	0.09	0.34
Italy	2.21	2.93	2.45	2.50	2.65	0.44	1.50
Mexiko	2.38	2.30	2.07	2.12	2.48	0.09	0.13
Norway	1.36	1.48	1.38	1.33	1.28	-0.08	-0.29
South Africa	1.36	1.44	1.39	1.57	1.50	0.14	1.02
South Korea	1.71	1.50	1.73	2.14	1.61	-0.10	-1.06
Spain	2.18	2.05	2.09	1.78	1.72	-0.46***	-2.71
Sweden	1.79	0.81	0.90	1.04	0.88	-0.92**	-2.17
Switzerland	1.30	1.52	1.01	0.97	1.57	0.27	1.28
Taiwan	1.78	1.76	1.65	1.73	1.76	-0.02	-0.54
$\mathbf{U}\mathbf{K}$	1.57	1.50	1.24	1.32	1.34	-0.23***	-2.65
USA	1.33	1.12	1.23	1.20	1.36	0.03	0.46

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pooled sample and the individual countries. Panel A reports the results of the Frazzini, Individual, Lai and Marmi factor sets whereas Panel B displays This table summarizes the recession coefficients of all factor set regressions. All regressions are panel fixed effect regressions (on the fund level) for the the factor set results of the regional factors. The dependent variable in all regressions is the TNA-weighed monthly fund return (excess over risk-free rate) as calculated in Section 3 for 16 countries. The independent variables are the Carhart (1997) factors (Market, SMB, HML, MOM) per country, from ECRI, OECD). The conditional specification displays the recession coefficients of the following regressions: (1) fund performance regressed on the Carhart Factors and the business cycle variable (dummy), (2) fund performance regressed on the Carhart Factors and the business cycle variables as well as their interactions (interactions). Furthermore, in subsamples of recession and non-recession periods, the fund returns are regressed on the Carhart Factors the respective factor dataset as listed in the table, and a business cycle variable (BC) based on the countries' respective business cycle measure (NBER, (recession, non-recession). Statistical significance at the ten, five and one-percent level is indicated by *,**, and ***, respectively.

							Panel A:	Panel A: Overview Results Factor Sets - Country Factors	Results Fa	actor Sets	- Country	/ Factors								
Factor Set	Factor Set BC-Variable	Specification Alpha	Alpha	Pooled	Canada	Denmark	France	Germany	Italy	Japan	Mexico Norway		South Africa South Korea	South Korea	Spain	Sweden	Switzerland Taiwan	Taiwan	UK	USA
Frazzini	NBER/ECRI	NBER/ECRI Conditional	BC-Dummy Interactions Non-Recession	-0.38 *** -0.34 *** -0.03 ***		-0.63*** -0.66*** 0.25***	-0.81*** -0.84*** 0.15***	-0.39^{***} -0.45^{***} 0.14^{***}	-1.04*** -1.1*** -0.04	-0.41*** -0.01 -0.23***		-0.79*** -0.8*** 0.6***			$\begin{array}{c} 0.01 \\ 0.32^{***} \\ -0.45^{***} \end{array}$	-1.6*** -1.34*** 0.51***	-1.44^{***} -1.14^{***} -0.01	- 1 '		-0.08*** 0.00 -0.04***
	OECD	Conditional	Recession BC-Dummy Interactions Non-Recession Recession	-0.38 *** -0.28 *** -0.23 *** -0.04 *** -0.38 ***	-0.53*** 0.08*** 0.01 -0.11*** -0.1***	-0.42*** -0.84*** -0.69*** 0.39*** -0.28***	-0.73*** -1.02*** -0.85*** 0.17*** -0.66***	-0.31*** -0.89*** -0.8*** 0.25*** -0.51***	-1.1^{***} -0.49^{***} -0.44^{***} -0.2^{***} -0.57^{***}	-0.21*** -0.22*** 0.21*** -0.32***		-0.21 *** -0.31 *** -0.17 *** 0.24 *** 0.07 ***			-0.1^{***} -0.22^{***} 0.01 -0.29^{***} -0.29^{***}	-1.05^{**} -0.84^{***} -0.87^{***} 0.59^{***} -0.31^{***}	-1.12*** -0.88*** -0.65*** -0.09***		0.35*** -0.79*** -0.51*** -0.34***	-0.03^{***} 0.01 0.02^{**} -0.04^{***} -0.03^{***}
Individual	Individual NBER/ECRI Conditional	Conditional	BC-Dummy Interactions Non-Recession	-0.16^{***} -0.11^{***} -0.01^{***}		0.31^{***} 0.21^{*} -0.12^{***}	0.15^{***} 0.02 -0.26^{***}	-0.06 0.03 -0.09***	-0.08* -0.07 -0.04***	-0.3*** -0.31*** 0.12***	-0.16 -0.16 0.12*	0.3^{***} 0.35^{***} 0.03	-0.63 *** -0.62 *** 0.23 ***	-0.03 -0.13*** 0.1***	0.09*** 0.06** 0.06**	-0.25^{**} -0.34^{***} 0.05	-0.07 -0.06 -0.08***			-0.09*** -0.01 -0.03***
	OECD	Conditional	Recession BC-Dummy Interactions Non-Recession Recession	-0.11 ** * -0.06 ** * -0.02 ** * -0.11 ** *	-0.35*** -0.01 -0.03 -0.1*** -0.13***	$\begin{array}{c} 0.08\\ 0.24^{***}\\ 0.22^{***}\\ -0.09^{***}\\ 0.14^{***}\end{array}$	-0.24*** -0.17*** -0.17*** -0.11*** -0.28***	-0.04 -0.24*** -0.3*** -0.03 -0.03	-0.07 *** 0.09*** 0.1*** 0.1*** 0.00	-0.18^{***} -0.01 -0.04^{**} 0.05^{***}	-0.06 -0.03 0.04 0.01	$\begin{array}{c} 0.36^{*} **\\ 0.39^{*} **\\ 0.33^{*} **\\ 0.33^{*} **\\ 0.03\\ 0.03\end{array}$	-0.39*** -0.28*** -0.33*** 0.2*** -0.14***	-0.04 -0.31 *** -0.31 *** 0.14 *** -0.2 ***	-0.09*** -0.05* -0.05* -0.15*** -0.2***	-0.33*** 0.01 -0.01 0.00 0.00	-0.16*** -0.09** -0.08*** -0.17***	- 0.76***	0.11^{***} - 0.11^{***} - 0.11^{***} - 0.05 - -0.01 - 0.02^{*} - -0.02^{*} - -100^{*} - -10	-0.03^{***} 0.00 -0.04^{***} -0.03^{***}
Lai	NBER/ECRI OECD	NBER/ECRI Conditional OECD Conditional	BC-Dummy Interactions Non-Recession Recession BC-Dummy Interactions Non-Recession Recession	-0.4** -0.48** 0.02*** -0.47*** -0.32*** 0.01** 0.01**	-0.64^{***} -0.63^{***} -0.01 -0.69^{***} 0.08^{***} 0.14^{***} -0.15^{***}	$\begin{array}{c} -0.88***\\ -0.18\\ 0.21***\\ 0.04\\ -0.78***\\ -0.12\\ 0.19***\\ 0.09**\end{array}$	-0.65*** -0.12** 0.2*** 0.07** -1.09*** -0.53*** 0.24***	0.03 0.19** -0.07** 0.13* -0.48*** -0.31*** 0.07	-1.15** -0.93*** 0.02 -0.83*** -0.37*** -0.03 -0.27***	-0.25*** -0.26*** -0.45*** -0.45*** 0.01 -0.3*** -0.3***	-0.61*** -0.98*** 0.26*** -0.71*** -0.9*** -1.24*** 0.36***	-1.27*** -0.99*** 0.7*** -0.29*** -0.64*** 0.62*** -0.01	-2.18 *** -2.27 *** 1.01 *** -1.34 *** -0.92 *** 0.5 *** 0.5 ***	-0.6*** -2.95*** 0.14*** -2.83*** 0.44*** 0.71*** 0.71*** 0.45***	-0.33*** 0.34*** -0.72*** -0.38*** -0.31*** 0.17*** -0.52***	-1.17*** -1.3*** 0.3*** -1.05*** -0.9*** 0.52*** 0.52***	$\begin{array}{c} -1.39^{***}\\ -0.44^{***}\\ 0.14^{***}\\ -1.29^{***}\\ -1.01^{***}\\ -0.23^{***}\\ -0.04\end{array}$	-0.14** -0.27*** 0.15*** -0.17*** 	-0.13*** 0.06 0.06** -0.17*** -0.93*** -0.52*** -0.48***	$\begin{array}{c} -0.08^{***} \\ 0.00 \\ -0.04^{***} \\ -0.03^{***} \\ 0.01 \\ 0.02^{**} \\ -0.04^{***} \end{array}$
Marmi	NBER/ECRI OECD	NBER/ECRI Conditional OECD Conditional	BC-Dummy Interactions Non-Recession Recession BC-Dummy	-0.114*** -0.11*** -0.08*** -0.19***			0.04 -0.1*** -0.27*** -0.36***	-0.39*** -0.62*** -0.05** -0.6*** 0.05						-0.34** -0.38*** -0.11*** -0.59***		-0.4*** -0.38** 0.12** -0.31*** 0.28**	-0.09* -0.01 -0.18*** -0.21***			-0.09*** 0.00 -0.04*** -0.03***
			Interactions Non-Recession Recession	-0.09*** -0.09*** -0.19***	-0.3^{***} -0.16^{***} -0.46^{***}		-0.07*** -0.3*** -0.36***	0.2^{***} - 0.32^{***} - 0.07^{**}	-0.07^{*} -0.14^{***} -0.19^{***}	-0.13^{**} -0.1^{**} -0.22^{***}				-0.35*** -0.05*** -0.4***		0.47^{***} -0.02 0.42^{***}	-0.07* -0.17*** -0.26***		0.2^{***} (-0.13 ^{***} -1 0.07 ^{***} -1	0.02^{***} - 0.04^{***} - 0.02^{***}

						'n	anel B: O	Panel B: Overview Results Factor Sets - Regional Factors	sults Fac	tor Sets -	Regional	Factors								
Factor Set	BC-Variable Specification Alpha	Specification	Alpha	Pooled Canada		Denmark	France	Germany	Italy	Japan	Mexico	Norway	South Africa	South Korea	\mathbf{S} pain	Sweden	Switzerland	Taiwan	UK	\mathbf{USA}
Regional	NBER/ECRI	Conditional	NBER/ECRI Conditional BC-Dummy	-0.28*** -0.44***	-0.44***	-1.19***	-0.34***			-0.44***	-0.76***	-1.11^{***}	-0.82***	-0.9***	-0.13***	-1.59***		-1.42***	0.29^{***}	-0.03**
I				-0.22***		-1.15^{***}	-0.36***	0.66^{***}	-0.56***	-0.17***	-0.9***	-1.14^{***}	-0.78***	-1.21^{***}	-0.13^{***}	-1.88***	-1.06^{***}	-1.63^{***}	0.3^{***}	-0.05***
			Non-Recession	0.00	0.15^{***}		-0.05***		-0.22^{***}	-0.21^{***}	1.16^{***}	1.08^{***}	1.08^{***}		-0.11^{***}	_		-0.61***	0.08^{***}	+**60.0-
				-0.23^{***}				0.81^{***}		-0.35***	0.22^{***}	-0.07	0.27^{***}	-0.71^{***}	-0.23^{***}		-0.81***	-2.25^{***}	0.46^{***}	-0.14^{***}
	OECD	Conditional	Conditional BC-Dummy	-0.22^{***}	-0.24^{***}			-0.58***	0.08	-0.19^{***}	-0.25	-0.6***	-0.57***		0.18^{***}	-1.12^{***}	-0.68***		-0.62^{***}	0.04^{***}
			Interactions	-0.2***	-0.15^{***}	-0.89***		-0.66***	0.13	0.13^{***}	-0.37**	-0.48^{***}	-0.48^{***}		0.28^{***}	-1.39^{***}	-0.66***		-0.75***	0.02
			Non-Recession	0.01^{**}	0.2^{***}	0.64^{***}			-0.61^{***}	-0.35^{***}	~	0.73^{***}	1.07^{***}			1.23^{***}	0.21^{***}		0.66^{***}	-0.1^{***}
			Recession	-0.21^{***}	0.04^{***}	-0.24^{***}		-0.28***		-0.22***	0.7^{***}	0.26^{***}	0.55^{***}		-0.02	-0.22***	-0.45^{***}		-0.07***	-0.1***
Borional	NREP /FCB1	Conditional		***℃	***00 0	- MOX -	***	***000	1 00***	***34 0	- OT**	1 0***	***℃ U	***08 U	0 20***	1 81**	1 28***	1 20***	*40 0	0.01
Tregrona .		CONTRICTORIAL		0.0	-0-24		1.10-	0.4.0		-010		-1.1- 		-0.07	-0.0-	10.1-				10.0-
(Momentum)				-0.33***	-0.21***	-1.51 ^{***}	-0.19 ^{***}	0.87 ***		-0.22***	-1.15 ^{***}	-0.8***	-0.95***	-1.37***	-0.13**	-2.12***	-1.49 ^{***}	-2.28***		-0.03***
				-0.02^{***}	0.1^{***}		0.05^{***}	0.19^{***}	-0.2***	-0.16^{***}	1.11^{***}	0.75^{***}	0.77^{***}	0.32^{***}	+60.0-	0.98^{***}	0.24^{***}	-0.45^{***}		-0.1^{***}
				-0.35^{***}	-0.08***		-0.18^{***}	1.06^{***}	-1.12^{***}	-0.34^{***}	-0.05	-0.05	-0.22**	-1.03^{***}		-1.15^{***}	-1.24^{***}	-2.77***	0.27^{***}	-0.13^{***}
	OECD	Conditional		-0.22***	-0.07***	-1.02^{***}			-0.35**		-0.4**	-0.65***	-0.05	-0.19^{***}		-1.66^{***}	-0.98***		-0.94^{***}	0.05^{***}
				-0.18^{***}	0.04			-0.57***	-0.25	0.12^{***}	-0.2	0.11	0.15^{***}	-0.09***	-0.06	-1.48***	-0.83***		-1.02^{***}	0.02^{**}
			e	-0.02^{***}	0.1^{***}	0.52^{***}				-0.35^{***}	1.12^{***}	0.21^{***}	0.57^{***}			0.71^{***}	0.15^{***}		0.61^{***}	-0.1^{***}
			Recession	-0.33***	0.13^{***}		-0.47***	-0.34^{***}		-0.21^{***}	0.89^{***}	0.33^{***}	0.69^{***}		-0.06^{***}	-0.76***	-0.65***		-0.4***	-0.1^{***}

Table A.2: Overview Factor Set Results (ctd.)

Table A.3: (Recession) Performance and Fund Flows

This table report results of univariate portfolio sorts based on unconditional performance in month t and future fund flows in month t + 1 as well as conditional performance in month t and future fund flows in month t + 1. for 16 countries and the pooled sample. The fund flows are calculated as described in Section 3. Statistical significance at the ten, five and one-percent level is indicated by *,**, and ***, respectively.

Panel A: Fund Flows sorted on Unconditional Performance

	Q1	$\mathbf{Q2}$	Q3	$\mathbf{Q4}$	$\mathbf{Q5}$	Q5-Q1	T-stat
ALL	0.016	0.021	0.018	0.021	0.031	0.014***	7.57
Canada	0.050	0.030	0.025	0.016	0.046	-0.004	-0.42
Denmark	0.017	0.007	0.001	-0.003	0.003	-0.014	-0.99
France	0.010	0.012	0.008	0.007	0.025	0.016	1.29
Germany	-0.001	0.006	-0.004	-0.002	0.021	0.022^{**}	2.07
Italy	0.012	0.005	0.006	0.002	-0.013	-0.026**	-2.15
Japan	0.008	-0.001	0.008	0.016	0.021	0.013^{**}	2.13
Mexiko	0.018	0.026	0.029	0.019	0.030	0.012	0.68
Norway	0.051	-0.004	0.010	0.022	0.033	-0.019	-0.64
South Africa	0.017	0.025	0.009	0.010	0.011	-0.005	-0.79
South Korea	0.011	0.017	0.024	0.031	0.052	0.041^{***}	5.17
\mathbf{Spain}	0.017	-0.003	-0.002	-0.006	0.010	-0.007	-0.67
Sweden	0.032	0.040	0.041	0.019	0.011	-0.021	-1.16
Switzerland	0.011	0.005	0.001	0.018	0.026	0.015	1.37
Taiwan	-0.013	-0.005	-0.006	0.000	-0.001	0.012^{***}	2.43
UK	-0.007	-0.005	0.001	0.001	0.023	0.031^{***}	3.95
\mathbf{USA}	0.026	0.030	0.022	0.025	0.035	0.009^{***}	3.28

Panel B: Fund Flows sorted on Conditional Performance

	Q1	$\mathbf{Q2}$	$\mathbf{Q3}$	$\mathbf{Q4}$	$\mathbf{Q5}$	Q5-Q1	T-stat
ALL	0.021	0.022	0.023	0.018	0.022	0.001	0.61
Canada	0.042	0.031	0.040	0.029	0.051	0.009	0.89
Denmark	0.011	0.008	-0.003	0.002	0.004	-0.008	-0.96
France	0.010	0.015	0.012	0.012	0.000	-0.010	-1.30
Germany	-0.001	0.007	0.003	0.007	-0.002	0.000	-0.02
Italy	0.002	-0.006	0.015	-0.003	0.011	0.009	0.64
Japan	0.019	0.013	0.002	-0.001	0.007	-0.012^{*}	-1.72
Mexiko	0.024	0.021	0.025	0.024	0.026	0.002	0.10
Norway	0.037	0.016	0.025	0.011	0.037	0.001	0.03
South Africa	0.009	0.021	0.008	0.014	0.023	0.014	1.53
South Korea	0.027	0.037	0.035	-0.017	0.021	-0.006	-0.67
\mathbf{Spain}	0.008	0.000	-0.007	-0.002	0.017	0.009	0.82
Sweden	0.027	0.031	0.022	0.014	0.040	0.013	0.70
$\mathbf{Switzerland}$	0.003	-0.002	0.005	0.029	0.026	0.023^{**}	2.01
Taiwan	-0.011	-0.002	-0.006	-0.004	-0.004	0.007	1.47
$\mathbf{U}\mathbf{K}$	-0.007	-0.005	-0.003	0.001	0.024	0.031^{***}	3.83
\mathbf{USA}	0.031	0.028	0.026	0.020	0.029	-0.002	-0.66

Table A.4: Summary Statistics: Hedge Funds Dataset

This table shows summary statistics of the TASS data sample of hedge funds and different hedge fund strategies in the time period from 1994 until 2012.

Strategy	Category	Funds	Average Monthly	Average	Average Net Assets
			Excess Return	Standard Deviation	(in USD)
1-10	All	15,332	0.13%	0.0363	7.41bn
1	Fixed Income Arbitrage /	623	0.25%	0.0275	4.72bn
	Convertible Arbitrage				
2	Dedicated Short Bias	51	-0.25%	0.0547	618m
3	Emerging Markets	885	0.32%	0.0523	$3.05\mathrm{bn}$
4	Equity Market Neutral	617	0.10%	0.0260	9.35bn
5	Event Driven	745	0.33%	0.0310	4.13bn
6	Fund of Funds	$5,\!613$	-0.10%	0.0270	947m
7	Global Macro	698	0.26%	0.0378	1.32bn
8	Long/Short Equity Hedge	3,387	0.27%	0.0429	$7.8\mathrm{bn}$
9	Managed Futures	860	0.16%	0.0513	4.94bn
10	Multi-Strategy	$1,\!853$	0.30%	0.0312	39.49bn

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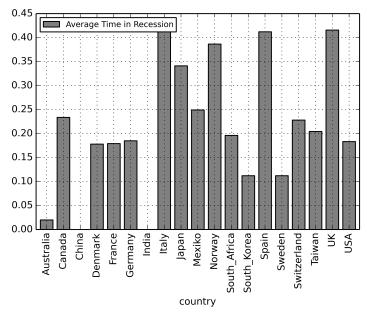
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Figures and Tables

Figure 1: Mutual Funds in Recessions (per Country)

Panel A: Times in Recessions. The figure displays the percentage of months a country spent in a recession during the sample period.



Panel B: Time Series of Recession Clustering. The figure displays the clustering of recessions in the sample period. The panel 'All Countries' shows the percentage of all sample countries that are in a recession. The panels 'North America', 'Europe', 'Asia', and 'Rest of the World' display the value of one if at least one country was in a recession in a certain month. North America includes: USA, Canada, Mexico; Europe includes: Denmark, France, Germany, Italy, Norway, Spain, Sweden, Switzerland, UK; Asia includes: Japan, South Korea; Rest of the World includes: South Africa, Australia.

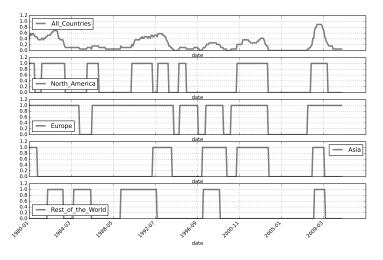


Figure 2: Average Returns of Funds per Country

The figure displays the average monthly returns (in %) of funds in the sample time period per country.

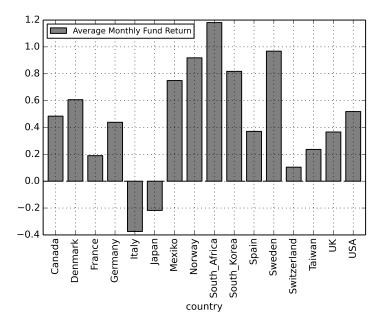
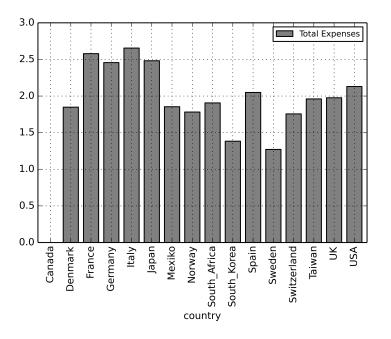


Figure 3: Total Expenses

The figure displays the total expenses (in % of a fund's TNA) in the observed time period per country. It is calculated as in Sirri and Tufano (1998) as "the expense ratio plus the up-frond load amortized over a seven-year holding period".



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This table provides summary statistics and reports the time horizon over which the mutual fund data and factor data is available. The number of funds gives the number of mutual funds (per fund) per country (after filtering, see Section 3). Australia, Austria, China and India drop out of the sample because the number of funds in the respective country or because the number of recession observations is not sufficient. The mean fund TNA displays the Regional-, Frazini-, Marmi- and Lai factors report the starting point of the factor data sets for the respective countries. All factor sets end in December 2010. average total net assets per fund. Fund Data (Start / End) shows the start / end time points of the mutual fund return data. The columns Individual-,

Country	Number of Funds (ID)	Number of Observations	Mean Fund TNA (in local currency)	Fund Data (Start)	Fund Data (End)	Individual Factors (Start)	Regional Fama- French Factors (Start)	Frazini Factors (Start)	Marmi Factors (Start)	Lai Factors (Start)
Australia		,		1980-01	2010-12	2001-01	1990-11	1990-07	1988-07	1981-07
Austria	ı			1986-07	2010-12	2001-02	1991-07	1991-07		1991-07
Canada	432	54,185	264m	1980-01	2010-12	2001-02	1990-11	1990-07	1990-07	1981-07
China	ı			2002-07	2010 - 12	2001-01	1993-01	ı	1998-07	1996-07
Denmark	39	5,765	566m	1980-01	2010-12	2001-02	1990-11	1990-07		1989-07
France		34,906	159m	1980-04	2010-12	2001-02	1990-11	1990-07	1988-07	1981-07
Germany	113	9,131	$311 \mathrm{m}$	1990-11	2010 - 12	2001-02	1990-11	1990-07	1988-07	1981-07
India	ı	1		1986-10	2010-12	10-7901	1993-01		1993-07	1993-07
Italy		5,566	$145 \mathrm{m}$	1984-11	2010-12	2001-02	1990-11	1990-07	1988-07	1988-07
Japan	811	104,495	$14.9 \mathrm{bn}$	1980-01	2010-12	2001-02	1994-11	1990-07	1988-07	1981-07
Mexico		3,262	314m	1980-01	2010 - 12	1996-06	1993-02			1993-07
Norway		7,049	774m	1981-03	2010-12	2001-02	1993-01	1990-07		1990-07
South Africa		13,826	408m	1980-01	2010 - 12	1994-06	1993-01			1993-01
South Korea	524	42,600	$69.4 \mathrm{bn}$	1996-02	2010-12	1997-01	1992-02		1992-07	1989-07
Spain		11,463	$53 \mathrm{m}$	1990-10	2010-12	2001-02	1990-11	1990-07		1989-07
Sweden	108	4,655	$2.1 \mathrm{bn}$	1985-01	2010-12	2001-02	1990-11	1990-07	1988-07	1988-07
Switzerland		13,675	$267 \mathrm{m}$	1980-01	2010 - 12	2001-02	1990-11	1990-07	1988-07	1988-07
Taiwan	182	20,553	$1.7\mathrm{bn}$	1986-02	2010-12	1997-01	1990-11			1994-07
UK	458	23,417	117m	1980-01	2010-12	2001-02	1990-11	1990-07	1988-07	1981-07
A SU	3.692	460523	553m	1980-01	2010-12	1980-01	1980-01	1990-07	1988-07	1981-07

Table 2: Fund Performance in Recessions - Pooled Results

This table shows the results of panel fixed effect regressions (on the fund level) of unconditional and conditional mutual fund performance for our pooled sample of 16 countries. The dependent variable is the TNA-weighed monthly fund return (excess over risk-free rate) as calculated in Section 3. The independent variables are the Carhart (1997) factors per country from the Lai factor dataset as described in Section 3, and a business cycle variable (BC) based on the countries respective business cycle measure (NBER, ECRI). Statistical significance at the ten, five and one percent level is indicated by *,**, and ***, respectively. Clustered robust standard errors (by fund) are in parentheses.

	(1)	(2)	(3)	(4)	(5)
	Unconditional	Performance	Recessions	No Recessions	Conditional Performance
	Performance	and BC	Subsample	Subsample	and Interactions
BC		-0.402***			-0.480***
		(0.00996)			(0.0106)
Market	0.789^{***}	0.783***	0.686***	0.798^{***}	0.797***
	(0.00307)	(0.00312)	(0.00414)	(0.00316)	(0.00314)
$Market \times BC$		· · · · ·	()		-0.110***
					(0.00300)
SMB	0.180***	0.181^{***}	0.0686***	0.186^{***}	0.187***
	(0.00484)	(0.00485)	(0.00693)	(0.00497)	(0.00497)
$SMB \times BC$		× ,	· · · · ·		-0.127***
					(0.00646)
HML	0.00918**	0.00721	-0.0177***	0.0138^{***}	0.0137***
	(0.00443)	(0.00442)	(0.00549)	(0.00469)	(0.00468)
$HML \times BC$			()		-0.0242***
					(0.00465)
MOM	-0.0218***	-0.0247***	-0.156***	0.0213***	0.0208***
	(0.00195)	(0.00197)	(0.00291)	(0.00261)	(0.00260)
$MOM \times BC$		· · · · ·	()		-0.172***
					(0.00379)
Constant	-0.0229***	0.0647^{***}	-0.466***	0.0220***	0.0211***
	(0.00239)	(0.00367)	(0.00725)	(0.00383)	(0.00371)
Observations	757,859	757,859	156,112	601,747	757,859
R-squared	0.668	0.669	0.725	0.629	0.673
Number of funds	7,321	7,321	7,287	7,298	7,321

Table 3: Fund Performance in Recessions - Individual Countries

country, obtained from the Lai factor dataset as described in Section 3, and a business cycle variable (BC) based on the countries respective business cycle measure (NBER, ECRI). In Panel A, fund performance is regressed on the Carhart factors and a business cycle variable. In Panel B, fund performance is monthly fund return (excess over risk-free rate) as calculated in Section 3 for 16 countries. The independent variables are the Carhart (1997) factors per This table summarizes results from panel fixed effect regressions (on the fund level) in 16 different countries. The dependent variable is the TNA-weighed regressed on the Carhart factors and the business cycle variable as well as their interactions. Statistical significance at the ten, five and one percent level is indicated by *, **, and ***, respectively. Clustered robust standard errors (by fund) are in parentheses.

	(1) Canada	(2) Denmark	(3) France	(4) Germany	(5) Italy	(6) Japan	(7) Mexiko	(8) Norway	(9) South_Africa	(10) South_Korea	(11) Spain	(12) Sweden	(13) Switzerland	(14) Taiwan	(15) UK	(16) USA
Market	0.532***	0.608***	0.573***	0.610***	0.578***	0.796***	0.592***	0.635***	0.231***	0.633***	0.588***	0.592^{***}	0.500***	0.888***	0.509^{***}	0.993***
SMB	(0.101^{***})	(0.001^{64})	(0.00339 -0.0339	(0.0230) -0.0230	(0.0157) -0.244***	(0.0576***	(0.0209) -0.111***	(0.147***	(0.000 - 0.480***	(0.0962^{***})	-0.0608***	(ncsnn:n) +8080:0-	-0.397***	(0.505***	(0.372***	0.176***
	(0.0124)	(0.0361)	(0.0300)	(0.0490)	(0.0413)	(0.0116)	(0.0304)	(0.0389)	(0.0277)	(0.00789)	(0.0180)	(0.0418)	(0.0299)	(0.0128)	(0.0220)	(0.00674)
HML	0.0624*** (0.00061)	-0.104***	-0.114^{***}	0.159^{***}	-0.0207	-0.0268*	0.157***	0.0755***	-0.338***	-0.0868***	-0.183*** (0.00725)	0.0307	0.00863	-0.243*** /0.00765\	-0.104*** (0.0136)	0.0673***
MOM	(TOGOO.0)	-0.153^{***}	-0.223^{***}	-0.139***	-0.103^{***}	-0.0430^{***}	-0.203^{***}	-0.0756***	0.0334^{***}	0.158***	-0.0509^{***}	-0.0523^{***}	-0.0772^{***}	0.154^{***}	-0.300***	0.00357*
i	(0.00452)	(0.0113)	(0.00070)	(0.0105)	(0.0260)	(0.00730)	(0.0395)	(0.00868)	(0.0103)	(0.00462)	(0.00829)	(0.00956)	(0.00869)	(0.00623)	(0.00925)	(0.00209)
BC	-0.638^{***} (0.0413)	-0.880*** (0.0998)	-0.655^{***}	0.0352 (0.0863)	-1.154^{***} (0.0732)	-0.247^{***} (0.0243)	-0.607^{***}	-1.266^{***} (0.0817)	-2.181^{***}	-0.603^{***} (0.0373)	-0.325^{**}	-1.209^{***} (0.188)	-1.391^{***} (0.0736)	-0.141^{**}	-0.130^{**}	-0.0849*** (0.00944)
Constant	(0.0127)	(0.0402)	(0.0153)	(0.0334)	(0.0451)	(0.00896)	(0.0580)	(0.0417)	(0.0210)	(0.0109)	(0.0218)	(0.0422)	(0.0274)	(0.0134^{***})	(0.0166)	(0.00325)
Observations R-squared Number of funds	49,799 0.627 432	4,971 0.702 39	31,756 0.571 299	8,095 0.713 113	4,920 0.755 75	96,332 0.569 811	2,727 0.700 49	6,046 0.769 61	$\begin{array}{c} 11,866\\ 0.553\\ 173\end{array}$	36,550 0.767 524	10,222 0.772 139	$3,743 \\ 0.776 \\ 108$	$\begin{array}{c} 11,809\\ 0.555\\ 169\end{array}$	$\begin{array}{c} 18,554 \\ 0.784 \\ 182 \end{array}$	$\begin{array}{c} 19,286 \\ 0.712 \\ 458 \end{array}$	$\begin{array}{c} 441,210\\ 0.756\\ 3,692\end{array}$
				1	Panel B: Mı	ttual Fund	Performan	ice and the	Business Cy	Panel B: Mutual Fund Performance and the Business Cycle with Interactions	actions					
	(1) Canada	(2) Denmark	(3) France	(4) Germany	(5) Italy	(6) Japan	(7) Mexiko	(8) Norway	(9) South_Africa	(10) South_Korea	(11) Spain	(12) Sweden	(13) Switzerland	(14) Taiwan	(15) UK	$^{(16)}_{ m USA}$
BC	-0.635***	-0.174	-0.123^{**}	0.189^{**}	-0.927***	-0.217***	-0.979***	-0.988***	-2.274***	-2.947***	0.344^{***}	-1.319***	-0.442***	-0.269***	0.0585	0.00451
Market	(0.0462) 0.531***	(0.117) 0.644***	(0.0494) 0.650***	(0.0909) 0.655***	(0.0500) 0.602***	(0.0264) 0.803***	(0.158) 0.575***	(0.0736) 0.502***	(1680.0)	(0.120) 0.658***	(0.0551***	(0.147) 0.605***	(0.0654) 0 401 ***	(1.001) 0.894***	(0.0409) 0.688***	(c0600.0) 0 986***
	(0.00823)	(0.00892)	(0.0109)	(0.0150)	(0.0184)	(0.00713)	(0.0212)	(0.0132)	(0.00834)	(0.00424)	(0.0314)	(0.0032)	(0.00728)	(0.00570)	(0.00826)	(0.00280)
Market*BC	0.00542 (0.00712)	-0.103^{***} (0.0161)	-0.404^{***} (0.0145)	-0.282^{***} (0.0239)	-0.251^{***} (0.0201)	-0.177^{***} (0.00922)	-0.127^{***} (0.0183)	0.0263^{*} (0.0138)	-0.0935^{***} (0.0110)	-0.677^{***} (0.0251)	-0.112^{***} (0.0284)	-0.158^{***} (0.0317)	-0.128^{***} (0.0204)	-0.0662^{***} (0.00884)	-0.383^{***} (0.00996)	0.0527^{***} (0.00273)
SMB	0.140^{***}	-0.350***	0.0452	0.00949	-0.210***	0.211^{***}	-0.0446	-0.128***	-0.448**	0.0990***	-0.178***	-0.0221	-0.492***	0.505***	-0.101***	0.180^{***}
SMB*BC	(0.0129) - 0.246^{***}	(0.0322) 0.282***	(0.0323) - 0.274^{***}	(0.0413) -0.0494	(0.0442) -0.173***	(0.0112) - 0.471 ***	(0.0318) - 0.487^{***}	(0.0240) - 0.158^{***}	(0.0250) -0.410***	(0.00764) - $0.877***$	(0.0270) -0.0331	(0.0459) - 0.354^{***}	(0.0362) 0.306^{***}	(0.0137) - 0.0456^{**}	(0.0260) -0.519***	(0.00685) - 0.0235^{***}
HML	(0.0211) 0.0993^{***}	(0.0454) -0.0731***	(0.0277) -0.133***	(0.0577) 0.142^{***}	(0.0473) - 0.0942^{***}	(0.0108) 0.0443^{***}	(0.0571) 0.198^{***}	(0.0334) 0.199^{***}	(0.0321) -0.336***	(0.0290) - 0.0733^{***}	(0.0244) - 0.188^{***}	(0.0503) 0.187^{***}	(0.0551) 0.133^{***}	(0.0176) -0.233***	(0.0257) - 0.137^{***}	(0.00482) 0.0945^{***}
HML*BC	(0.0102) -0.251***	(0.0149) -0 202***	(0.0162) 0.0303	(0.0192) 0.0986***	(0.0257) 0.0884***	(0.0145) -0.105***	(0.0250) 0.200***	(0.0118) -0.438***	(0.0162) -0.0507	(0.00763) -0.164***	(0.0113) 0.134***	(0.0325) -1 112***	(0.0257) -0.288***	(0.00771) -0.0683***	(0.0233) -0.0533**	(0.00682) -0.143***
	(0.0147)		(0.0247)	(0.0299)	(0.0298)	(0.0155)	(0.0453)	(0.0266)	(0.0400)	(0.0297)	(0.0136)	(0.0716)	(0.0271)	(0.00974)	(0.0232)	(0.00446)
MUM	-0.0277^{***} (0.00540)	-0.112^{***} (0.0105)	-0.132^{***} (0.00950)	-0.0333^{***} (0.0114)	(0.0338)	(0.0821^{***})	(0.0647^{*})	(0.0159)	(0.018^{***})	(0.00499)	(0.0692)	-0.0508^{***} (0.0118)	0.268^{***} (0.0123)	(0.00651)	(0.0297)	(0.00287)
MOM*BC	-0.00484	-0.174^{***}	-0.582***	-0.383***	-0.573***	-0.407***	-0.757***	-0.505***	-0.239***	-1.041***	-0.709***	-0.219^{***}	-0.910***	-0.105^{**}	-0.597***	-0.0162***
Constant	(0.0135)	(0.0384) (0.0384)	(0.0137^{***}) (0.0132)	(0.0323)	(0.0365)	(0.00973)	(0.0538)	(0.0365) (0.0365)	(0.0364)	(0.0108)	(0.0532) (0.0532)	(0.0408)	(0.0276) (0.0276)	(0.0129) (0.0129)	(0.0294) (0.0294)	(0.00347)
Observations B conneed	49,799	4,971	31,756 0.601	8,095 0.796	4,920 0.785	96,332 0 507	2,727 0.793	6,046 0 803	11,866	36,550 0 803	10,222	3,743 0.708	11,809	18,554 0.785	19,286 0 727	441,210
Number of funds		39	299	113	75	811	49	0.000 61	173	0.000 524	130	0.790 108	169	199	10.101	101.0

Table 4: USA - Comparison of Results

This table summarizes results from panel fixed effect regressions (on the fund level) in the USA. In all regressions the excess fund return is regressed on the four U.S. Carhart factors from Kenneth Frenchs website and a business cycle variable (BC) based on the NBER business cycle measure as well as their interactions. Regression (1) is directly taken from the results of Glode (2008). Regressions (2) and (4) are replications of the Glode (2008) results with data from the CRSP database in the time period from 1980 - 2005 and from 1980 - 2010, respectively. Regressions (3) and (5) repeat the same regressions with Morningstar mutual fund data. Statistical significance at the ten, five and one percent level is indicated by *,**, and ***, respectively. Clustered robust standard errors (by fund) are in parentheses.

	(1)	(2)	(3)	(4)	(5)
	Glode (2008)	(2)	Morningstar	CRSP	Morningstar
	1980-2005	1980-2005	1980-2005	1980-2010	1980-2010
BC	0.414***	0.000973***	0.0704***	-9.12e-05	0.00451
	(0.078)	(0.000198)	(0.0159)	(0.000113)	(0.00905)
Market	0.994***	0.985***	1.000***	0.974***	0.986***
	(0.004)	(0.00436)	(0.00310)	(0.00390)	(0.00280)
Market*BC	0.028	-0.0340***	-0.0621***	0.0470***	0.0527***
	(0.020)	(0.00851)	(0.00872)	(0.00321)	(0.00273)
SMB	0.207***	0.187***	0.187***	0.181***	0.180***
	(0.009)	(0.00767)	(0.00724)	(0.00734)	(0.00685)
SMB*BC	-0.114***	-0.0318***	-0.0366***	-0.0381***	-0.0235***
	(0.017)	(0.00744)	(0.00647)	(0.00570)	(0.00482)
HML	0.050***	0.0867***	0.126***	0.0623***	0.0945***
	(0.009)	(0.00953)	(0.00802)	(0.00836)	(0.00682)
HML*BC	-0.034	-0.0950***	-0.0745***	-0.147***	-0.143***
	(0.032)	(0.00853)	(0.00664)	(0.00587)	(0.00446)
MOM	0.037***	0.0106***	0.0112***	0.0140***	0.0162***
	(0.005)	(0.00350)	(0.00325)	(0.00313)	(0.00287)
MOM*BC	-0.076***	-0.0682***	-0.0944***	-0.0250***	-0.0162***
	(0.016)	(0.0108)	(0.00990)	(0.00335)	(0.00308)
Constant	-0.367***	-0.000672***	-0.0755***	-0.000384***	-0.0424***
	(0.019)	(5.54e-05)	(0.00520)	(3.98e-05)	(0.00347)
Observations	82.081	273.632	285.765	393.124	441,210
R-squared	0.74	0.661	0.680	0.723	0.757
Number of funds	3,260	2.444	3.128	2.678	3,692

Table 5: Cross-Sectional Implications

This table displays the results of univariate portfolio sorts based on average monthly unconditional alphas according to the four-factor Carhart (1997) model. We display equal-weighted averages of unconditional alphas (% per month), fund expense ratios (% of TNA), total fund fees (% of TNA) and recession alphas (% per month) for each of the quintile portfolios. Panel A displays the results based on the pooled worldwide sample and Panel B shows the results for individual countries. Canada and Japan are missing from the list of countries, as there not enough observations for fund fees. Statistical significance at the ten, five and one percent level is indicated by *,**, and ***, respectively. T-statistics are displayed in parentheses.

Panel A: Pooled Sample									
Qunitile	Unconditional Alpha	Expenses	Total Fee	Recession Alpha					
	(% per month)	(% of TNA)	(% of TNA)	(% per month)					
Q1	-0.63	1.80	2.39	-0.54					
Q2	-0.19	1.46	2.20	-0.17					
Q3	-0.04	1.22	2.00	-0.04					
Q4	0.10	1.32	1.98	0.10					
Q5	0.53	1.47	1.89	0.52					
Q5 - Q1	1.16***	-0.33***	-0.49***	1.07***					
	(97.36)	(-3.38)	(-3.00)	(6.07)					

	Panel B: Individual Countries									
Qunitile	Unconditional Alpha	Expenses	Total Fee	Recession Alpha						
	(% per month)	(% of TNA)	(% of TNA)	(% per month)						
Denmark	0.95***	-0.34	-0.33	1.19***						
Q5 - Q1	(8.84)	(-0.63)	(-0.59)	(5.39)						
France	1.15***	-0.62	-0.67*	1.34***						
Q5 - Q1	(21.58)	(-1.56)	(-1.73)	(7.27)						
Germany	1.32***	-0.03	0.00	1.46***						
Q5 - Q1	(12.95)	(-0.12)	(-0.01)	(7.51)						
Italy	0.91***	0.63	0.50	0.46***						
Q5 - Q1	(11.35)	(1.66)	(1.22)	(4.70)						
Mexico	0.95***	1.46**	1.46**	0.53***						
Q5 - Q1	(9.31)	(2.32)	(2.32)	(2.61)						
Norway	0.73***	0.13	0.16	0.74***						
Q5 - Q1	(8.29)	(0.50)	(0.55)	(5.27)						
South Africa	1.49***	-0.01	0.19	0.52***						
Q5 - Q1	(30.29)	(-0.08)	(1.30)	(5.15)						
South Korea	1.07***	-0.68***	-0.01	1.11***						
Q5 - Q1	(31.48)	(-8.35)	(-0.13)	(6.10)						
Spain	0.97***	-0.51***	-0.44***	0.63***						
Q5 - Q1	(13.51)	(-2.59)	(-2.37)	(5.15)						
Sweden	1.73***	0.01	0.31	1.00***						
Q5 - Q1	(21.19)	(0.01)	(0.57)	(6.77)						
Switzerland	0.91***	-0.43*	-0.54**	0.76***						
Q5 - Q1	(13.90)	(-1.91)	(-2.16)	(7.18)						
Taiwan	1.25***	-0.05	-0.02	1.21***						
Q5 - Q1	(13.52)	(-0.82)	(-0.30)	(7.92)						
UK	1.43***	-0.26***	-0.30***	1.56***						
Q5 - Q1	(29.13)	(-2.64)	(-2.49)	(7.16)						
USA	0.73***	-0.46***	-0.46*	0.68***						
Q5 - Q1	(58.37)	(-4.35)	(-1.88)	(8.77)						

Table 6: Risk Factors and Business Cycles

This table summarizes results from various panel fixed effects regressions (on the fund level). In all regressions the excess fund return is regressed on the Carhart (1997) factors and a business cycle variable (BC) based on the countries respective business cycle measure (NBER, ECRI, OECD) as well as their interactions. Regressions (1) and (6) are based on the Carhart (1997) factors of the Lai factor dataset, whereas regressions (2), (3), (4) and (5) are based on the Frazzini, Marmi, individual and regional factor datasets as described in Section 3. Regressions (1)-(5) use the NBER and ECRI business cycle measure, whereas regression (6) uses the OECD business cycle measure. Statistical significance at the ten, five and one percent level is indicated by *,**, and ***, respectively. Clustered robust standard errors (by fund) are in parentheses.

	(1)	(2)	(2)	(4)	(5)	(6)
	(1) Lai	(2) Enograini	(3) Marmi	(4) Individual	(5) Docional	(6) OECD
		Frazzini Frazzini			Regional	
	Factor-Set	Factor-Set	Factor-Set	Factor-Set	Factor-Set	Recession
BC	-0.480***	-0.340***	-0.113***	-0.109^{***}	-0.333***	-0.283***
	(0.0106)	(0.0102)	(0.00944)	(0.00864)	(0.00969)	(0.00923)
Market	0.797^{***}	0.854^{***}	0.972^{***}	0.945^{***}	0.851^{***}	0.787^{***}
	(0.00314)	(0.00336)	(0.00247)	(0.00216)	(0.00280)	(0.00319)
Market*BC	-0.110***	-0.0921***	-0.00490*	0.0263^{***}	-0.0862***	-0.0359***
	(0.00299)	(0.00295)	(0.00281)	(0.00213)	(0.00259)	(0.00192)
SMB	0.187^{***}	0.189^{***}	0.151^{***}	0.199^{***}	0.198^{***}	0.185^{***}
	(0.00497)	(0.00583)	(0.00474)	(0.00499)	(0.00574)	(0.00517)
SMB^*BC	-0.127^{***}	-0.168^{***}	-0.0438***	-0.0209***	-0.119^{***}	-0.0413***
	(0.00646)	(0.00644)	(0.00518)	(0.00472)	(0.00700)	(0.00337)
HML	0.0137^{***}	0.0417^{***}	0.0792^{***}	0.0507^{***}	0.0470^{***}	0.0426^{***}
	(0.00468)	(0.00488)	(0.00474)	(0.00508)	(0.00528)	(0.00419)
HML*BC	-0.0242***	-0.100***	-0.0475***	-0.0715***	-0.182^{***}	-0.00907**
	(0.00465)	(0.00378)	(0.00445)	(0.00435)	(0.00487)	(0.00425)
MOM	0.0208***	0.00658^{***}	0.0467^{***}		0.0318***	0.0467***
	(0.00260)	(0.00212)	(0.00232)		(0.00259)	(0.00309)
MOM*BC	-0.172***	-0.139***	-0.0240***		-0.152***	-0.124***
	(0.00379)	(0.00305)	(0.00289)		(0.00375)	(0.00334)
Constant	0.0211***	-0.0299***	-0.0835***	-0.00544^{**}	-0.0238***	0.0373***
	(0.00371)	(0.00352)	(0.00310)	(0.00260)	(0.00330)	(0.00467)
Observations	$757,\!859$	686,410	691,281	705,744	$738,\!459$	754,803
R-squared	0.673	0.685	0.760	0.782	0.639	0.668
Number of fund_id	7,321	6,393	$6,\!678$	$7,\!321$	7,321	$7,\!524$

Table 7: Fund Style

This tablerecession performance for different fund styles shows the (small/large/mid/value/growth/income/other). All regressions are panel fixed effects regressions (on the fund level) with a specification as outlined in equation (2). The dependent variable is the TNA-weighed monthly fund return (excess over risk-free rate) as calculated in Section 3, pooled over 16 countries. The independent variables are the Carhart (1997) factors per country, calculated as described in Section 3, and a business cycle variable (BC) based on the countries respective business cycle measures (NBER, ECRI). Statistical significance at the ten, five and one percent level is indicated by *,**, and ***, respectively. Clustered robust standard errors (by fund) are in parentheses.

	(1)	(2)	(3)	(4)	(5)	(7)	(8)
	small	large	mid	value	growth	income	other
DC	0.001***	0 510***	0.000***	0.00.1***	0 1 / 1 * * *	0 501***	0 101***
BC	-0.221***	-0.513***	-0.398***	-0.204***	-0.141***	-0.524***	-0.461***
	(0.0330)	(0.0125)	(0.0317)	(0.0284)	(0.0187)	(0.0453)	(0.0265)
Market	0.876***	0.773***	0.866***	0.860***	0.947***	0.664***	0.624***
	(0.00793)	(0.00395)	(0.00876)	(0.00840)	(0.00595)	(0.0194)	(0.00642)
Market*BC	-0.0272***	-0.113***	-0.0316***	-0.0288***	-0.0267***	-0.135***	-0.229***
	(0.00780)	(0.00387)	(0.00828)	(0.00760)	(0.00541)	(0.0144)	(0.00773)
SMB	0.594^{***}	0.0117^{***}	0.377^{***}	0.154^{***}	0.261^{***}	-0.0355***	-0.191***
	(0.00903)	(0.00374)	(0.00912)	(0.0112)	(0.00919)	(0.00823)	(0.0154)
SMB*BC	-0.100***	-0.146^{***}	-0.138***	-0.00325	-0.153^{***}	-0.0687***	-0.196^{***}
	(0.0145)	(0.00724)	(0.0149)	(0.0135)	(0.0113)	(0.0231)	(0.0167)
HML	0.0315^{**}	-0.0136**	-0.0440***	0.369^{***}	-0.128***	0.162^{***}	-0.0479^{***}
	(0.0144)	(0.00535)	(0.0126)	(0.00887)	(0.00779)	(0.0140)	(0.00884)
HML*BC	-0.0187*	-0.0107**	-0.0899***	-0.117***	-0.104***	0.00406	-0.107***
	(0.0104)	(0.00546)	(0.0104)	(0.0108)	(0.00682)	(0.0162)	(0.0162)
MOM	0.0581***	0.0158***	0.0623***	-0.0782***	0.0830***	-0.0643***	0.0130
	(0.00617)	(0.00310)	(0.00711)	(0.00407)	(0.00417)	(0.00750)	(0.00791)
MOM*BC	-0.143***	-0.167***	-0.167***	-0.000467	-0.145***	-0.0600***	-0.375***
	(0.00824)	(0.00513)	(0.00951)	(0.00632)	(0.00562)	(0.0116)	(0.00920)
Constant	0.0176*	0.00290	0.106***	-0.0137	-0.0423***	0.0611***	0.0475***
	(0.00970)	(0.00363)	(0.00838)	(0.00985)	(0.00574)	(0.0176)	(0.0113)
						. ,	
Observations	$137,\!458$	$396,\!677$	120,102	118,480	261,320	31,220	$65,\!642$
R-squared	0.700	0.725	0.699	0.762	0.718	0.683	0.729
Number of fund_id	1,325	3,773	1,182	1,085	$2,\!174$	325	796

Table 8: Fund Flows

This table shows dependent portfolio double sorts based on unconditional and conditional Carhart (1997) four factor alphas of mutual funds based on the pooled sample in month t. Fund flows in month t + 1 for the 25 portfolios are calculated as described in Section 3. Statistical significance at the ten, five and one percent level is indicated by *,**, and ***, respectively. T-statistics are displayed in parentheses.

Portfolio Double Sorts									
Conditional	Q1	$\mathbf{Q2}$	Q3	Q 4	$\mathbf{Q5}$	Q5-Q1	Average		
Unconditional									
$\mathbf{Q1}$	1.74%	0.92%	2.11%	1.03%	0.49%	$-1.26\%^{***}$	-		
						(-2.58)			
$\mathbf{Q2}$	1.82%	2.66%	2.37%	1.83%	1.20%	-0.62%*	-		
_						(-1.87)			
$\mathbf{Q3}$	1.95%	2.40%	1.92%	1.84%	1.18%	-0.78%**	-		
<u></u>	1 0107	0.001	0.0407	1.050	1.05%	(-1.97)			
$\mathbf{Q4}$	1.81%	2.60%	2.84%	1.95%	1.37%	-0.44%	-		
05	2.000	0.0007	0 5 707	0.0707	4 1007	(-0.98)			
$\mathbf{Q5}$	2.98%	2.02%	2.57%	2.27%	4.19%	1.22%***	-		
						(3.07)			
Q5-Q1	$1.23\%^{***}$	$1.11\%^{***}$	0.46%	$1.25\%^{**}$	$3.70\%^{***}$	-	1.55%		
	(4.34)	(3.36)	(1.16)	(2.32)	(6.61)				
Average	-	-	-	-	-	-0.38%	-		

Table 9: Tracking Errors and Aggregate Funds Flows

Panel A of this table reports the results of equal-weighted averages of mutual funds' tracking errors in recessions and non-recessions. Panel B reports aggregate fund flows into mutual funds in recessions and non-recessions. In each panel we also report differences with corresponding t-statistics. Statistical significance at the ten, five and one percent level is indicated by *,**, and ***, respectively.

	I allel A	.: Tracking Error	. 5	
Country	Recession	Non-Recession	Difference	T-statistics
	Tracking Error	Tracking Error		
All	4.16	3.41	0.74	20.89
Canada	5.85	4.30	1.55	16.41
Denmark	4.89	3.70	1.19	7.08
France	5.45	4.50	0.96	8.45
Germany	5.57	4.22	1.35	6.48
Italy	4.23	3.94	0.29	2.18
Japan	4.09	3.91	0.19	2.70
Mexiko	6.78	4.58	2.20	5.59
Norway	4.51	4.55	-0.04	-0.27
South_Africa	9.18	6.26	2.92	21.62
South_Korea	9.10	5.35	3.75	41.08
Spain	5.15	4.07	1.09	8.08
Sweden	5.76	4.36	1.40	7.95
Switzerland	5.46	4.00	1.46	6.59
Taiwan	5.09	4.84	0.25	2.39
UK	4.46	3.74	0.73	9.71
USA	2.61	2.45	0.17	4.83

Panel A: Tracking Errors

Panel B: Aggregate Fund Flows

Country	Recession	Non-Recession	Difference	T-statistics
	Aggregate Flows	Aggregate Flows		
All	-0.09%	1.61%	-1.70%***	-5.25
Canada	0.43%	2.10%	$-1.67\%^{**}$	-2.14
Denmark	-0.38%	1.07%	$-1.45\%^{***}$	-4.13
France	-0.64%	1.37%	$-2.01\%^{***}$	-5.14
Germany	0.04%	0.76%	-0.72%	-1.45
Italy	-1.09%	0.92%	$-2.01\%^{***}$	-4.19
Japan	0.09%	0.94%	-0.85%**	-2.04
Mexiko	-0.45%	1.68%	-2.13%**	-2.31
Norway	0.76%	2.24%	-1.48%***	-5.01
South_Africa	-1.28%	1.03%	$-2.31\%^{***}$	-5.31
South_Korea	1.68%	1.65%	0.03%	0.05
Spain	-0.56%	0.83%	$-1.39\%^{***}$	-5.62
Sweden	-0.18%	0.76%	-0.92%	-0.52
Switzerland	-0.73%	0.93%	$-1.66\%^{***}$	-6.12
Taiwan	-0.43%	0.30%	-0.73%	-0.65
UK	0.03%	0.95%	$-0.92\%^{***}$	-6.89
USA	0.26%	1.79%	-1.53%***	-6.33

Table 10: Explaining Mutual Fund Underperformance: Portfolio Sorts

This table reports the results of univariate portfolio sorts. We sort funds in quintiles according to their tracking error in recessions. Then we report the average Carhart (1997) four factor alpha for each of the quintiles. Statistical significance at the ten, five and one percent level is indicated by *,**, and ***, respectively.

	$\mathbf{Q1}$	$\mathbf{Q2}$	$\mathbf{Q3}$	$\mathbf{Q4}$	$\mathbf{Q5}$	Q5-Q1	Q5-Q1 t-value
Canada	-0.21	-0.22	-0.33	-0.19	-0.11	0.10	0.56
Denmark	-0.14	0.01	-0.29	-1.03	-1.35	-1.21***	-6.78
France	-0.05	-0.69	-0.75	-0.77	-0.71	-0.66***	-4.11
Germany	-0.31	-0.61	-1.01	-0.65	-0.34	-0.03	-0.45
Italy	-0.04	-0.38	-0.43	-0.35	-0.41	-0.37***	-4.11
Japan	0.06	-0.56	-0.58	-0.51	-0.76	-0.82***	-6.90
Mexiko	-0.33	-0.25	-0.33	-0.20	0.14	0.47^{*}	1.89
Norway	-0.45	-0.32	-0.25	-0.36	-0.37	0.08	0.99
South Africa	-0.42	-0.53	-0.44	-0.47	-1.21	-0.79***	-5.48
South Korea	-0.47	-0.53	-1.31	-1.38	-1.16	-0.69***	-3.17
Spain	-0.27	-0.75	-0.14	-0.17	-0.16	0.11	1.21
Sweden	-0.76	-0.95	-0.78	-0.85	-0.50	0.26^{***}	3.67
Switzerland	-0.26	-0.41	-0.54	-0.62	-0.41	-0.15^{**}	-2.34
Taiwan	-0.42	-0.27	-0.17	-0.28	-0.37	0.05	0.55
UK	0.21	0.37	0.07	-0.21	-0.51	-0.72***	-5.89
USA	0.34	0.32	-0.11	-0.34	-0.42	-0.76***	-7.81
ALL	0.12	-0.21	0.43	-0.34	-0.61	-0.73***	-5.74
ALL, Flows $< 0\%$	0.09	-0.13	0.32	-0.27	-1.25	-1.34***	-4.89
ALL, Flows $< -1\%$	0.08	-0.17	0.51	-0.39	-1.49	-1.57***	-4.12
ALL, Flows $< -2\%$	0.19	-0.11	0.23	-0.14	-1.76	-1.95***	-3.00

Table 11: Explaining Mutual Fund Underperformance: Regression Analysis

This table reports the results of country *i*'s average mutual fund recession performance in month t on country *i*'s average mutual fund recession tracking error (TE), average aggregate recession fund flows (FF) and different country characteristics measured in month t. As country characteristics we use GDP per capita, the number of listed domestic stocks, mutual fund assets per GPD per capita, the number of mutual fund companies per country, and a country's home bias to our model. Statistical significance at the ten, five and one percent level is indicated by *,**, and ***, respectively. Clustered robust standard errors (by country) are in parentheses.

	(1)	(2)	(3)	(4)
	Recession	Recession	Recession	Recession
	Performance	Performance	Performance	Performance
TE	-0.1659**		-0.1761**	-0.1969**
	(0.0847)		(0.0769)	(0.0581)
\mathbf{FF}	. ,	0.2623^{***}	0.2864^{*}	0.2398^{*}
		(0.0921)	(0.1497)	(0.1234)
GDP				0.053
				(0.099)
No. of Stocks				-0.0003
				(0.0007)
Mutual Fund Assets				-0.0002
				(0.0005)
No. of MF Companies				0.0008*
				(0.00045)
Home Bias				-0.3452*
				(0.1967)
Constant	0.1969	-0.6690***	0.3063	0.2567
	(0.4817)	(0.1470)	(0.4452)	(0.4152)
Observations	1097	1097	1097	1097
R-squared	0.163	0.094	0.297	0.356

Table 12: Hedge Funds

Panel A: This table shows the results of unconditional and conditional hedge fund performance measurement in panel fixed effects regressions (on the fund level) for the USA based on the seven-factor model by Fung and Hsieh (2004) and the NBER, ECRI business cycle variable (BC). The dependent variable is the monthly fund return (excess over risk-free rate) from the TASS database for the sample of all hedge funds. Statistical significance at the ten, five and one-percent level is indicated by *,**, and ***, respectively. Clustered robust standard errors (by fund) are in parentheses.

	(1)	(2)	(3)	(4)	(5)
	Unconditional	Performance	Recessions	No Recessions	Conditional Performance
	Performance	and BC	Subsample	Subsample	and Interactions
BC		-0.00571***			-0.00362***
DC		(0.000146)			(0.000147)
Market	0.190***	0.184***	0.172***	0.217***	0.215***
	(0.00292)	(0.00311)	(0.00358)	(0.00425)	(0.00421)
Market*BC	()	()	()	()	-0.0445***
					(0.00452)
SMB	0.0512***	0.0603***	-0.105***	0.100***	0.0992***
	(0.00302)	(0.00328)	(0.00522)	(0.00360)	(0.00361)
SMB*BC		~ /			-0.205***
					(0.00593)
PTFSBD	-0.00137***	0.00204***	0.0149^{***}	0.00367^{***}	0.00367***
	(0.000432)	(0.000476)	(0.00122)	(0.000581)	(0.000583)
PTFSBD*BC					0.0113***
					(0.00114)
PTFSFX	0.00473^{***}	0.00106^{***}	-0.0111***	0.00523^{***}	0.00549^{***}
	(0.000269)	(0.000297)	(0.000700)	(0.000376)	(0.000376)
PTFSFX*BC					-0.0157^{***}
					(0.000785)
PTFSCOM	0.00670^{***}	0.00762^{***}	0.00541^{***}	0.00904^{***}	0.00955^{***}
	(0.000395)	(0.000454)	(0.00107)	(0.000520)	(0.000519)
PTFSCOM*BC					-0.00558***
					(0.00110)
Treasurychange	0.00347^{***}	0.00376^{***}	0.0127^{***}	-0.00731***	-0.00669***
	(0.000283)	(0.000310)	(0.000617)	(0.000349)	(0.000350)
Treasurychange*BC					0.0194***
					(0.000671)
Moodieschange	-0.0173***	-0.0159***	-0.0134***	-0.0230***	-0.0219***
	(0.000309)	(0.000323)	(0.000345)	(0.000519)	(0.000521)
Moodieschange*BC					0.00817***
					(0.000573)
Constant	0.000710***	0.00266***	-0.00208***	0.00218***	0.00205***
	(1.87e-05)	(4.20e-05)	(9.74e-05)	(4.46e-05)	(4.93e-05)
Observations	623,228	484,754	121,818	362,936	484,754
R-squared	0.117	0.124	0.152	0.109	0.136
Number of funds	15,332	15,058	9,398	13,830	15,058

Table 12: Hedge Funds (continued)

Panel B: This table shows the results of conditional hedge fund performance measurement in panel fixed effects regressions (on the fund level) for the USA based on the seven-factor model by Fung and Hsieh (2004) and the NBER, ECRI business cycle variable (BC). The dependent variable is the monthly fund return (excess over risk-free rate) from the TASS database for the strategy subsamples of (1) Convertible/Fixed Income Arbitrage, (2) Dedicated Short Bias, (3) Emerging Markets, (4) Equity Market Neutral, (5) Event Driven, (6) Fund of Funds, (7) Global Macro, (8) Long/Short Equity Hedge, (9) Managed Futures, (10) Multi Strategy hedge funds. Statistical significance at the ten, five and one-percent level is indicated by *,**, and ***, respectively. Clustered robust standard errors (by fund) are in parentheses.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
VARIABLES	Strategy 1	Strategy 2	Strategy 3	Strategy 4	Strategy 5	Strategy 6	Strategy 7	Strategy 8	Strategy 9	Strategy 10
DC	0.000.400	0.00000	0 0051 (***	0 001 =0***	0 00 10 5***	0.00500***	0.00114	0 0001 0***	0 000 - 1 * * *	0 00000***
BC	0.000433	0.00220	-0.00514***	-0.00178***	-0.00497***	-0.00522***	0.00114	-0.00216***	-0.00271***	-0.00360***
NC 1 /	(0.000798) 0.0534^{***}	(0.00262) -0.758***	(0.000728) 0.380***	(0.000552) 0.0266**	(0.000524) 0.159^{***}	(0.000167) 0.172^{***}	(0.00103) 0.143^{***}	(0.000326) 0.350***	(0.000696) 0.204***	(0.000672) 0.127***
Market	(0.0534)	(0.0802)	(0.380) (0.0166)	(0.0200^{-1})	$(0.159^{-1.1})$ (0.00873)	(0.172^{+++}) (0.00479)	(0.0143)	$(0.350^{-1.1})$	(0.204) (0.0166)	(0.127) (0.00996)
Market [*] BC	(0.0105) 0.0751^{***}	0.326***	(0.0100) 0.0384^*	0.0299**	0.0263**	-0.0400***	-0.0207	-0.0616***	-0.300***	-0.00232
Market DC	(0.0751) (0.0136)	(0.0817)	(0.0384) (0.0219)	(0.0299) (0.0143)	(0.0203)	(0.00499)	(0.0207)	(0.0105)	(0.0254)	(0.00232) (0.0112)
SMB	0.0440***	-0.233***	0.112***	0.0378***	0.0876***	0.0867***	-0.00496	0.190***	-0.0183	0.0400***
SMD	(0.00828)	(0.0463)	(0.0112)	(0.0578)	(0.00909)	(0.0007)	(0.0162)	(0.130) (0.00850)	(0.0133)	(0.0400)
SMB*BC	-0.176***	-0.0922	-0.254***	-0.125***	-0.130***	-0.232***	-0.141***	-0.269***	-0.0243	-0.127***
DMD DO	(0.0220)	(0.119)	(0.0301)	(0.0221)	(0.0178)	(0.00662)	(0.0331)	(0.0139)	(0.0310)	(0.0193)
PTFSFD	-0.0107***	-0.0283***	-0.00979***	-0.00751***	-0.0119***	0.00166**	0.00420	0.00606***	0.0508***	0.00142
	(0.00163)	(0.00896)	(0.00223)	(0.00199)	(0.00159)	(0.000779)	(0.00291)	(0.00109)	(0.00339)	(0.00172)
PTFSFD*BC	-0.00597	-0.0299	0.0403***	0.00557	-0.000917	0.00752***	0.00298	0.0157***	0.00248	0.00573
	(0.00429)	(0.0236)	(0.00517)	(0.00543)	(0.00377)	(0.00134)	(0.00648)	(0.00266)	(0.00654)	(0.00354)
PTFSFX	-0.00256***	-0.00581	0.000421	0.00240**	0.000900	0.00481***	0.0202***	0.00352***	0.0339***	-0.00261**
	(0.000853)	(0.00512)	(0.00175)	(0.00106)	(0.000943)	(0.000517)	(0.00247)	(0.000636)	(0.00249)	(0.00114)
PTFSFX*BC	-0.000185	0.0362**	-0.0316***	-0.00856***	0.00761***	-0.0116***	-0.0349***	-0.0132***	-0.0742***	-0.00139
	(0.00286)	(0.0144)	(0.00388)	(0.00275)	(0.00259)	(0.000982)	(0.00507)	(0.00169)	(0.00528)	(0.00258)
PTFSCOM	0.000443	0.00722	0.00848^{***}	0.00781^{***}	0.000602	0.00921^{***}	0.0144^{***}	0.00785^{***}	0.0389^{***}	0.00761^{***}
	(0.00106)	(0.00768)	(0.00273)	(0.00141)	(0.00100)	(0.000589)	(0.00253)	(0.00104)	(0.00369)	(0.00151)
PTFSCOM*BC	-0.0277***	-0.0133	0.0144^{***}	-0.0110***	-0.0164^{***}	-0.0164^{***}	0.00950	0.00132	0.0640^{***}	-0.00941^{***}
	(0.00466)	(0.0212)	(0.00542)	(0.00397)	(0.00347)	(0.00134)	(0.00745)	(0.00237)	(0.00729)	(0.00359)
Treasurychange	-0.0156^{***}	0.000775	-0.00738***	-0.00416^{***}	-0.00805***	-0.00430***	-0.0107***	-0.00496***	-0.0151^{***}	-0.00592***
	(0.00155)	(0.00484)	(0.00175)	(0.00104)	(0.000903)	(0.000433)	(0.00245)	(0.000644)	(0.00243)	(0.00110)
Treasurychange*BC	-0.000143	-0.0412***	0.0302***	0.00593**	0.0186***	0.0143***	0.0179***	0.0227***	0.0345***	0.0138***
	(0.00311)	(0.0127)	(0.00344)	(0.00268)	(0.00229)	(0.000746)	(0.00407)	(0.00156)	(0.00338)	(0.00204)
Moodieschange	-0.0404***	-0.0124	-0.0248***	-0.0115***	-0.0373***	-0.0203***	-0.0262***	-0.0174***	-0.0294***	-0.0177***
M P 1 *DG	(0.00254)	(0.00814)	(0.00263)	(0.00189)	(0.00185)	(0.000598)	(0.00357)	(0.00113)	(0.00289)	(0.00160)
Moodieschange*BC	0.0107***	-0.00249	0.0112***	-0.000691	0.0148***	0.00113*	0.0299***	0.00885***	0.0403***	0.00376**
G	(0.00270)	(0.00969)	(0.00302)	(0.00225)	(0.00210)	(0.000659)	(0.00425)	(0.00125)	(0.00330)	(0.00171)
Constant	0.00214***	0.000245	0.00542***	0.00134***	0.00371***	0.000757***	0.00193***	0.00255***	0.00279***	0.00300***
	(0.000147)	(0.000593)	(0.000232)	(0.000154)	(0.000133)	(7.35e-05)	(0.000291)	(9.19e-05)	(0.000173)	(0.000203)
Observations	24,222	2,066	30,460	20,747	31,064	163,708	17.853	129,150	28,880	36,604
R-squared	0.145	0.424	0.229	0.047	0.215	0.251	0.039	0.198	0.082	0.094
Number of funds	622	50	870	606	736	5,506	686	3,345	830	1,807